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**"WEAK CONVERGENCE OF SAMPLE COVARIANCE
MATRICES TO STOCHASTIC INTEGRALS VIA
MARTINGALE APPROXIMATIONS"**

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WEAK CONVERGENCE OF SAMPLE COVARIANCE MATRICES
TO STOCHASTIC INTEGRALS VIA MARTINGALE APPROXIMATIONS

by

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0. ABSTRACT

Under general conditions the sample covariance matrix of a vector martingale and its differences converges weakly to the matrix stochastic integral $\int_0^1 B dB'$, where B is vector Brownian motion. For strictly stationary and ergodic sequences, rather than martingale differences, a similar result obtains. In this case, the limit is $\int_0^1 B dB' + \Lambda$ and involves a constant matrix, Λ , of bias terms whose magnitude depends on the serial correlation properties of the sequence. This note gives a simple proof of the result using martingale approximations.

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1. INTRODUCTION

There has recently been a good deal of interest in time series regressions that involve integrated processes. The theory makes extensive use of weak convergence methods in general, and multivariate invariance principles in particular. Some recent papers dealing with this topic are [1, 3-16]. Much of the theory involves weak convergence of sample covariance matrices to matrix stochastic integrals of the form $\int_0^1 B dB' + \Lambda$, where B is vector Brownian motion and Λ is a constant matrix of bias terms. This result is so important that it almost rivals the invariance principle in terms of its significance for applications.

To fix ideas, let $(x_t)_0^\infty$ be an n -vector time series generated by

$$x_t = x_{t-1} + u_t ; \quad t = 1, 2, \dots \quad (1)$$

where x_0 is any random vector (including a constant) and $(u_t)_{-\infty}^\infty$ is a zero mean, strictly stationary and ergodic sequence with continuous spectral density $f_{uu}(\lambda)$. Define $X_T(r) = T^{-1/2} \sum_1^{[Tr]} u_j$. Then, as shown in [8, 12], under quite general conditions as $T \rightarrow \infty$ we have:

$$X_T(r) \Rightarrow B(r) = BM(\Omega) \quad (2)$$

with

$$\Omega = 2\pi f_{uu}(0) = \Sigma + \Lambda + \Lambda'$$

$$\Sigma = E(u_0 u_0') , \quad \Lambda = \sum_{k=1}^{\infty} E(u_0 u_k')$$

and

$$T^{-1} \sum_{t=1}^T x_{t-1} u'_t \Rightarrow \int_0^1 B dB' + \Lambda . \quad (3)$$

Here, we use the symbol " \Rightarrow " to signify weak convergence as $T \rightarrow \infty$, " \equiv " to signify equality in distribution and " $BM(\Omega)$ " to denote Brownian motion with covariance matrix Ω .

The proof of (3) that is given in [8] is lengthy and uses the concept of a near-integrated process [9, 10]. A more direct proof of the result seems desirable. When (u_t) forms a square integrable martingale difference sequence with respect to the natural filtration of σ -fields $F_t = \sigma(u_t, u_{t-1}, \dots)$, then $\Lambda = 0$ and (3) has been proved recently by direct methods in [1]. In particular we have:

LEMMA (Chan and Wei)

If (u_t, F_t) is a martingale difference sequence, if

$$E(u'_t u_t | F_{t-1}) \leq c \text{ a.s.} \quad (4)$$

for some constant $c > 0$, and if (2) holds then

$$T^{-1} \sum_{t=1}^T x_{t-1} u'_t \Rightarrow \int_0^1 B dB' .$$

The purpose of the present note is to show how (3) may be obtained quite simply when $\Lambda \neq 0$ by using this Lemma and a martingale approximation to the process u_t . The approach we follow is inspired by the use of martingale approximations in central limit theory for stationary processes. The reader is referred to [2, Ch. 5] for an excellent exposition of the approach.

2. MAIN RESULT AND PROOF

It will be convenient to let u_t in (1) be generated by the linear process

$$u_t = \sum_{j=-\infty}^{\infty} B_j e_{t-j}, \quad \sum_{j=-\infty}^{\infty} \|B_j\| < \infty \quad (5)$$

where the sequence of random vectors $(e_t)_{-\infty}^{\infty}$ is iid(0, Δ) with $\Delta > 0$ and where $\|B_j\| = \max_k (\sum_l |b_{jkl}|)$ with $B_j = (b_{jkl})$. This includes all stationary and invertible ARMA processes, for instance, and is therefore of wide applicability. The process u_t defined by (5) is strictly stationary and ergodic and has continuous spectral density given by

$$f_{uu}(\lambda) = (1/2\pi) (\sum_j B_j e^{ij\lambda}) \Delta (\sum_j B_j e^{ij\lambda})^* .$$

In addition to the absolute summability of $\{B_j\}$ in (5), we will use the following condition (based on (5.37) of Hall and Heyde [2])

$$\sum_{k=1}^{\infty} [\|\sum_{j=k}^{\infty} B_j\| + \|\sum_{j=k}^{\infty} B_{-j}\|] < \infty, \quad (6)$$

which is again satisfied by all stationary and invertible ARMA models.

Our main result is as follows:

THEOREM. *If (x_t) is generated by (1) and (u_t) satisfies (5) and (6) then (3) holds.*

PROOF. Under the stated conditions, we note first that the multivariate invariance principle (2) applies. When $n = 1$ this follows directly from Theorem 5.5 of Hall and Heyde [2, p. 141 and p. 146]. For $n > 1$, the

result may again be deduced from this theorem by applying the argument of Theorem 2.1 of [7].

The remainder of the theorem is based on a martingale approximation of u_t . The construction is achieved in Theorems 5.4 and 5.5 of Hall and Heyde [2]. We let $M_k = \sigma(e_j, j \leq k)$ and define

$$Y_0 = \sum_{\ell=-\infty}^{\infty} [E(u_\ell | M_0) - E(u_\ell | M_{-1})] = (\sum_{\ell=-\infty}^{\infty} B_\ell) e_0$$

$$Z_0 = \sum_{k=0}^{\infty} E(u_k | M_{-1}) = \sum_{k=-\infty}^{-1} (u_k - E(u_k | M_{-1})) .$$

Setting $Y_k = U^k Y_0$ and $Z_k = U^k Z_0$, where U is the temporal displacement operator, we observe that (Y_k, M_k) is a martingale difference sequence whose differences Y_k are strictly stationary, ergodic and square integrable with covariance matrix $\Omega = (\sum_j B_j) \Delta (\sum_j B_j') = 2\pi f_{uu}(0)$. The process (Z_k) is also strictly stationary, ergodic and square integrable. With this construction we have $u_0 = Y_0 + Z_0 - Z_1$ and thus

$$u_t = Y_t + Z_t - Z_{t+1} .$$

Note that $x_t = \sum_1^t u_j + x_0$ and writing $P_k = \sum_1^k Y_j$ we obtain

$$\begin{aligned} T^{-1} \sum_1^T x_{k-1} u_k' &= T^{-1} \sum_1^T (P_{k-1} + Z_1 - Z_k + x_0) (Y_k + Z_k - Z_{k+1})' \\ &= T^{-1} \sum_1^T P_{k-1} Y_k' + T^{-1} \sum_1^T (Z_1 - Z_k) Y_k' \\ &\quad + T^{-1} \sum_1^T P_{k-1} (Z_k - Z_{k+1})' + T^{-1} \sum_1^T (Z_1 - Z_k) (Z_k - Z_{k+1})' + o_p(1) . \end{aligned} \quad (7)$$

Now by ergodicity we have:

$$T^{-1} \Sigma_1^T Z_1 Y'_k \rightarrow 0 \quad \text{a.s.}$$

$$T^{-1} \Sigma_1^T Z_k Y'_k \rightarrow E(Z_0 Y'_0) \quad \text{a.s.}$$

$$T^{-1} \Sigma_1^T (Z_1 - Z_k)(Z_k - Z_{k+1})' \rightarrow -E(Z_0(Z_0 - Z_1)') \quad \text{a.s.}$$

and

$$\begin{aligned} T^{-1} \Sigma_1^T P_{k-1} (Z_k - Z_{k+1})' &= T^{-1} \Sigma_1^T P_{k-1} Z'_k - [T^{-1} \Sigma_1^T P_k Z'_{k+1} - T^{-1} \Sigma_1^T Y_k Z'_{k+1}] \\ &= T^{-1} P_0 Z_1 - T^{-1} P_{T+1} Z_{T+1} + T^{-1} \Sigma_1^T Y_k Z'_{k+1} \\ &\rightarrow E(Y_0 Z'_1) \quad \text{a.s.} \end{aligned}$$

Moreover, by the lemma

$$T^{-1} \Sigma_1^T P_{k-1} Y'_k = \int_0^1 B dB'$$

where $B(r) = BM(\Omega)$. We deduce that

$$T^{-1} \Sigma_1^T x_{t-1} u'_t = \int_0^1 B dB' + K$$

where

$$\begin{aligned} K &= E(Y_0 Z'_1) - E(Z_0(Z_0 - Z_1)') - E(Z_0 Y'_0) \\ &= E(Y_0 Z'_1) - E(Z_0 u'_0) . \end{aligned}$$

Now

$$Z_1 = \Sigma_{k=0}^{\infty} E(u_{k+1} | M_0) - \Sigma_{k=-\infty}^{-1} (u_{k+1} - E(u_{k+1} | M_0))$$

and

$$\begin{aligned} E(Y_0 Z_1') &= -\sum_{k=-\infty}^{-1} E(Y_0 u'_{k+1}) + \sum_{k=-\infty}^{\infty} E(Y_0 E(u_{k+1} | M_0)) \\ &= \sum_{k=0}^{\infty} E(Y_0 u'_{k+1}) , \end{aligned}$$

since Y_0 is M_0 -measurable. Next

$$\begin{aligned} E(Z_0 u'_0) &= -\sum_{k=-\infty}^{-1} E(u_k u'_0) + \sum_{k=-\infty}^{\infty} E(E(u_k | M_{-1}) u'_0) \\ &= -\sum_{k=1}^{\infty} E(u_0 u'_k) + \sum_{k=-\infty}^{\infty} E(E(u_k | M_{-1}) u'_0) . \end{aligned}$$

Hence,

$$\begin{aligned} K &= \Lambda + \sum_{k=0}^{\infty} E(Y_0 u'_{k+1}) - \sum_{k=-\infty}^{\infty} E(E(u_k | M_{-1}) u'_0) \\ &= \Lambda + \sum_{j=1}^{\infty} E(Y_{-j} u'_0) - \sum_{k=-\infty}^{\infty} E(E(u_k | M_{-1}) u'_0) . \end{aligned} \quad (8)$$

Now

$$Y_{-j} = \sum_{\ell=-\infty}^{\infty} (E(u_{\ell} | M_{-j}) - E(u_{\ell} | M_{-j-1}))$$

and

$$\begin{aligned} \sum_{j=1}^{\infty} E(Y_{-j} u'_0) &= \sum_{\ell=-\infty}^{\infty} (\sum_{j=1}^{\infty} [E(E(u_{\ell} | M_{-j}) u'_0) - E(u_{\ell} | M_{-j-1}) u'_0]) \\ &= \sum_{\ell=-\infty}^{\infty} E(E(u_{\ell} | M_{-1}) u'_0) \end{aligned} \quad (9)$$

since $E(u_{\ell} | M_{-\infty}) = 0$ a.s. We deduce from (8) and (9) that $K = \Lambda$ and (3) follows immediately.

3. SOME REMARKS ON APPLICATIONS

Limit theorems involving stochastic integrals such as (3) seem to be of widespread importance. They have many applications in econometrics and arise frequently in time series regressions with integrated processes and autoregressions with unit roots. Many examples are provided in the papers [3-6, 12-15]. In addition, as indicated in other recent work [11], it seems likely that a general asymptotic theory for optimization estimators can be developed that uses limit theorems such as (3). With some extensions, this theory can accommodate limits to stochastic integrals that are taken with respect to more general continuous parameter martingales. Some of the interesting possibilities for such extensions are explored in Section 4 of [11].

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