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ASYMPTOTICS FOR NONLINEAR TRANSFORMATIONS OF INTEGRATED  
TIME SERIES

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# Asymptotics for Nonlinear Transformations of Integrated Time Series<sup>1</sup>

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## Abstract

An asymptotic theory for stochastic processes generated from nonlinear transformations of nonstationary integrated time series is developed. Various nonlinear functions of integrated series such as ARIMA time series are studied, and the asymptotic distributions of sample moments of such functions are obtained and analyzed. The transformations considered in the paper include a variety of functions that are used in practical nonlinear statistical analysis. It is shown that their asymptotic theory is quite different from that of integrated processes and stationary time series. When the transformation function is exponentially explosive, for instance, the convergence rate of sample functions is path-dependent. In particular, the convergence rate depends not only the size of the sample, but also on the realized sample path. Some brief applications of these asymptotics are given to illustrate the effects of nonlinearly transformed integrated processes on regression. The methods developed in the paper are useful in a project of greater scope concerned with the development of a general theory of nonlinear regression for nonstationary time series.

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*Key words and phrases:* Additive functionals of Brownian motion, Brownian motion, integrated process, local time, nonlinear transformation, occupation time, regression asymptotics.

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# 1 Introduction

Nonstationary time series arising from autoregressive models with roots on the unit circle have been an intensive subject of recent research. The asymptotic behaviour of regression statistics based on integrated time series (those for which one or more of the autoregressive roots are unity) has received the most attention and a fairly complete theory is now available for linear time series regressions. The resulting limit theory forms the basis of much ongoing empirical econometric work, especially on the subject of unit root testing and cointegration modeling. The main elements of this limit theory as it is needed for linear regression was reviewed in Phillips (1988), and a recent overview of the asymptotic statistical theory on which some of the literature draws was given in Jeganathan (1995).

As in other regression contexts, linear models can be restrictive and they eliminate many interesting cases of practical importance where there are nonlinear responses to covariates. However, extensions of the existing limit theory for integrated processes to nonlinear models is not straightforward. This is because nonlinear functions of integrated processes often depend on fine-grain details of the underlying process, most especially the sojourn time that the process spends in the vicinity of certain points. These details need to be dealt with in the development of a limit theory for the sample functions that arise in regression.

The present paper seeks to provide some tools that will be useful in the analysis of time series regressions that involve nonlinear functions of integrated processes. Various nonlinear functions that commonly arise in practical nonlinear statistical analysis are studied. The results show that the limit theory can be very different from that for simple linear and polynomial functions of integrated processes. The case of exponential functions is especially interesting, because here the sojourn time that the process depends in the neighbourhood of its extrema determines the asymptotic behavior of the sample function. In consequence, the convergence rate of sample moments of exponential functions of the process is path dependent and relies on extreme sample path realizations of the time series.

## 2. Assumptions and Preliminary Results

We consider a time series  $\{x_t\}$  generated by

$$x_t = x_{t-1} + w_t, \tag{1}$$

where the error  $w_t$  follows the linear process

$$w_t = \varphi(L)\varepsilon_t = \sum_{k=0}^{\infty} \varphi_k \varepsilon_{t-k}, \tag{2}$$

in which  $\{\varepsilon_t\}$  is a sequence of i.i.d. random variables with mean zero, and for which  $\varphi(1) \neq 0$ . The system (1) is initialized at  $t = 0$  with  $x_0 = O_p(1)$ . One of the following two assumptions will be made throughout the paper.

**2.1 Assumption**  $\sum_{k=0}^{\infty} k^{1/2} |\varphi_k| < \infty$  and  $\mathbf{E} \varepsilon_t^2 < \infty$ .

**2.2 Assumption** (a)  $\sum_{k=0}^{\infty} k |\varphi_k| < \infty$  and  $\mathbf{E} |\varepsilon_t|^p < \infty$  for some  $p > 2$ .

(b) *The distribution of  $\varepsilon_t$  is absolutely continuous with respect to the Lebesgue measure, and has characteristic function  $\phi(t)$  for which  $\lim_{t \rightarrow \infty} t^r \phi(t) = 0$  for some  $r > 0$ .*

For simplicity, assume  $\varphi(1) = 1$  and  $\mathbf{E} \varepsilon_t^2 = 1$ . Other values simply have a scaling effect in the subsequent analysis.

Construct the stochastic process

$$W_n^0(r) = \frac{1}{\sqrt{n}} \sum_{t=1}^{[nr]} w_t,$$

which takes values in  $D[0, 1]$ , the set of cadlag functions on the interval  $[0, 1]$ . Phillips and Solo (1992) show that Assumption 2.1 is sufficient to ensure that  $W_n^0$  converges weakly to a standard linear Brownian motion  $W$  on  $[0, 1]$ . In our context, it is more convenient to endow  $D[0, 1]$  with the uniform topology rather than the usual Skorohod topology [see Billingsley (1968), pp. 150–152]. It then follows from the so-called Skorohod representation theorem [e.g., Pollard(1984), pp. 71–72] that there exists  $W_n$  such that  $W_n =_d W_n^0$  in  $D[0, 1]$  and  $W_n \rightarrow_{\text{a.s.}} W$  uniformly on  $[0, 1]$ . Akonom (1993) gives a specific rate of convergence under Assumption 2.2(a) using strong approximation methods.

**2.3 Lemma** *Let  $n \rightarrow \infty$ .*

(a) *If Assumption 2.1 holds, then  $\sup_{0 \leq r \leq 1} |W_n(r) - W(r)| = o(1)$  a.s.*

(b) *If Assumption 2.2(a) holds, then  $\sup_{0 \leq r \leq 1} |W_n(r) - W(r)| = o_p(n^{-(p-2)/2p})$ .*

Our development relies on the local time  $L(t, s)$  of the Brownian motion  $W(t)$  at  $s$ .  $L(t, s)$  is a jointly continuous stochastic process for which the following important formula applies (e.g., Chung and Williams, 1990).

**2.4 Lemma** (Occupation Times Formula) *Let  $T$  be locally integrable. Then*

$$\int_0^t T(W(r)) dr = \int_{-\infty}^{\infty} T(s) L(t, s) ds$$

*for all  $t \in \mathcal{R}$ .*

The local time  $L(t, s)$  has an interpretation as a spatial ( $s$ ) occupation density for the Brownian motion  $W$ . From the continuity of  $L(t, \cdot)$ , Lemma 2.4 can be applied with  $T(x) = 1 \{|x - s| < \varepsilon\}$  to give

$$L(t, s) = \lim_{\varepsilon \rightarrow 0} \frac{1}{2\varepsilon} \int_0^t 1 \{|W(r) - s| < \varepsilon\} dr, \quad (3)$$

the representation which explains why  $L(\cdot, s)$  is called the local time of  $W$  at  $s$ .

We define

$$N_n(\nu_n; a, b) = \int_0^1 1\{a \leq \nu_n W_n(r) \leq b\} dr \quad (4)$$

and similarly

$$N(\nu_n; a, b) = \int_0^1 1\{a \leq \nu_n W(r) \leq b\} dr, \quad (5)$$

where  $a$  and  $b$  are nonrandom constants, and  $\nu_n > 0$  for all  $n$ . The following useful result is due to Akonom (1993).

**2.5 Lemma** *Let Assumption 2.2 hold. Then as  $n \rightarrow \infty$*

(a)  $\mathbf{E}(N_n(\nu_n; 0, \delta) - N_n(\nu_n; k\delta, (k+1)\delta))^2 \leq c \frac{\delta}{n\nu_n} \left(1 + \frac{k\delta^2 n \log n}{\nu_n^2}\right)$  for some constant  $c$ , and

(b)  $N_n(\nu_n; 0, \pi_n) = N(\nu_n; 0, \pi_n) + o_p\left(n^{-(2p-1)/3p+\varepsilon}\right)$  for  $\pi_n \geq \nu_n n^{-2(p+1)/3p}$  and any  $\varepsilon > 0$ .

It follows from (3) that  $(\nu_n/\pi_n)N(\nu_n; 0, \pi_n) \rightarrow_{a.s.} L(1, 0)$  as  $n \rightarrow \infty$ . And from Lemma 2.5(b),  $(\nu_n/\pi_n)N_n(\nu_n; 0, \pi_n) = L(1, 0) + o_p(1)$  for  $\pi_n \geq \nu_n n^{-(2p-1)/3p+\varepsilon}$  with some  $\varepsilon > 0$ . In this sense, an appropriately defined  $N_n$  approximates  $L$  for large  $n$ . Also  $nN_n(\nu_n; a, b)$  is the number of visits of the process  $\nu_n W_n(r)$  to the interval  $[a, b]$ .

### 3. Functions of Normalized Integrated Processes

We start by investigating the asymptotic behavior of functions of normalized integrated processes. Such functions sometimes arise in models formulated with nonlinear functions of standardized partial sums of stationary time series. Let  $T$  be a measurable transformation in  $\mathcal{R}$ . We will consider *regular* transformations  $T$  defined as follows.

**3.1 Definition** *A transformation  $T$  is said to be regular if and only if, on every compact set  $C$ , there exist  $\underline{T}_\varepsilon, \overline{T}_\varepsilon$  and  $\delta_\varepsilon > 0$  for each  $\varepsilon > 0$  satisfying*

$$\underline{T}_\varepsilon(x) \leq T(y) \leq \overline{T}_\varepsilon(x) \quad (6)$$

for all  $x, y \in C$  such that  $|x - y| < \delta_\varepsilon$ , and

$$\int_C (\overline{T}_\varepsilon - \underline{T}_\varepsilon)(x) dx \rightarrow 0 \quad (7)$$

as  $\varepsilon \rightarrow 0$

The class of regular transformations includes locally bounded monotone functions and continuous functions. For a locally bounded monotone increasing function, for instance, set  $\underline{T}_\varepsilon(x) = T(x - \varepsilon), \overline{T}_\varepsilon(x) = T(x + \varepsilon)$  and  $\delta_\varepsilon = \varepsilon$ . Likewise, we set  $\underline{T}_\varepsilon(x) = T(x) - \varepsilon$  and  $\overline{T}_\varepsilon(x) = T(x) + \varepsilon$  for a continuous function with the usual  $\delta_\varepsilon$  for the  $\varepsilon$ - $\delta$  formulation of uniform continuity. It is easy to see that conditions (6) and (7) are satisfied for such choices. They work for any compact set. It is also clear that finite sums of locally bounded monotone functions (and hence functions which are locally of bounded variation) and piecewise continuous functions are regular.

**3.2 Theorem** *Let Assumption 2.1 hold. If  $T$  is regular, then*

$$\frac{1}{n} \sum_{t=1}^n T\left(\frac{x_t}{\sqrt{n}}\right) \xrightarrow{d} \int_0^1 T(W(r))dr$$

as  $n \rightarrow \infty$ .

### 3.3 Remarks

(a) Any regular transformation  $T$  is locally integrable. The local integrability of  $T$  guarantees that the limiting distribution is well defined. Indeed,  $T$  is locally integrable if and only if

$$\Pr \left\{ \int_0^t T(W(r))dr \text{ exists for all } t \right\} = 1$$

[see, e.g., Karatzas and Shreve (1988), Proposition 6.27, p. 216]. We need a stronger condition to ensure that the limiting distribution is invariant across different data generating processes.

(b) Given a transformation  $T$  on  $\mathcal{R}$ , we define a functional  $\Pi_T$  on  $D[0, 1]$  given by

$$\Pi_T : f \mapsto \int_0^1 T(f(r)) dr$$

For  $T$  defining a continuous  $\Pi_T$  on  $D[0, 1]$ , the result in Theorem 3.2 follows directly from the continuous mapping theorem [e.g., Billingsley (1968), Theorem 5.1, p. 30]. Uniformly continuous  $T$  generate such a functional. If  $T$  is continuous, but not uniformly continuous, the corresponding  $\Pi_T$  is assured of being continuous only on  $C[0, 1]$ , a subset of  $D[0, 1]$ . But the continuous mapping theorem still applies, since  $C[0, 1]$  is of Wiener measure one. Indeed, the proof of Theorem 3.2 shows that, for any regular  $T$ ,  $\Pi_T$  is continuous on a subset of  $D[0, 1]$  with Wiener measure one.

(c) The functions

$$T(x) = \log |x| \quad \text{and} \quad T(x) = |x|^\kappa \quad \text{for} \quad -1 < \kappa < 0 \quad (8)$$

are locally integrable and therefore  $\int_0^1 T(W(r)) dr$  is well defined for such functions. However, they are *not* regular and Theorem 3.2 does not apply.

To deal with such functions we may proceed as follows. Let  $T$  be locally integrable with a pole or logarithmic type of discontinuity at a certain point, say, zero. Define

$$\begin{aligned} T_n(x) &= T(x)1\{|x| \geq c_n\} \\ &+ T(c_n)1\{0 < x < c_n\} + T(-c_n)1\{-c_n < x < 0\} \end{aligned} \quad (9)$$

Similar modifications can be made for transformations with discontinuities at points other than zero.

**3.4 Theorem** *Let  $T$  be locally integrable. Suppose for a sequence  $\{c_n\}$  such that  $c_n \rightarrow 0$  and  $c_n \geq n^{-2(p+1)/3p}$ ,*

$$|T(x) - T(y)| \leq \nu(c_n)|x - y|$$

with  $\nu(c_n) = O(n^{(p-2)/2p})$  for all  $x, y \in \{z \mid |z| \geq c_n\}$ , and  $T(\pm c_n) = O(n^{(2p-1)/3p+\varepsilon})$  for some  $\varepsilon > 0$ . If Assumption 2.2 holds, then

$$\frac{1}{n} \sum_{t=1}^n T_n \left( \frac{x_t}{\sqrt{n}} \right) \xrightarrow{d} \int_0^1 T(W(r)) dr$$

as  $n \rightarrow \infty$ .

### 3.5 Remarks

(a) The conditions in Theorem 3.4 require that the function  $T$  be Lipschitz continuous on  $\{x : |x| \geq c_n\}$ . Also, the value of the function  $T(\pm c_n)$  around the discontinuity point and the Lipschitz constant  $\nu(c_n)$  may not grow too quickly with  $n$ .

(b) For the logarithmic function  $T(x) = \log|x|$ , the conditions in Theorem 3.4 are satisfied with  $c_n = n^{-\delta}$  for any  $\delta$  such that  $0 < \delta \leq (p-2)/2p$ . For the reciprocal function  $T(x) = |x|^\kappa$  with  $-1 < \kappa < 0$ , one may choose  $c_n = n^{-\delta}$  for  $0 < \delta < (p-2)/2p(1-\kappa)$  to show that the result in Theorem 3.4 is applicable.

(c) For any fixed  $n$ ,  $T$  and  $T_n$  are identical over any finite set of nonzero points, if we take  $c_n$  to be smaller than the minimum of their moduli. Therefore, if  $\{x_t\}$  is driven by an error process whose underlying distribution is of the continuous type specified in Assumption 2.2(b), then  $T$  and  $T_n$  are practically indistinguishable in finite samples.

## 4. Additive Functionals of Brownian Motion

The asymptotic behavior of functions of unnormalized integrated processes can be quite different from the results in the previous section. In particular, the asymptotics depend in a more critical way on the properties of the functions involved. To illustrate the dependencies that arise, we first investigate the asymptotic behavior of additive functionals of Brownian motion given by

$$\int_0^{\lambda t} T(W(r)) dr$$

as  $\lambda \rightarrow \infty$ . The results from this section will be applicable in the statistical analysis of the data that are continuously recorded from Brownian motion, or in the development of the asymptotics when the sampling frequency, as well as the time span of the data, increases. Applications of this type occur in econometrics, especially with financial data (e.g. Shiller and Perron, 1986, and Phillips, 1987). More directly, the limit behavior of these functionals sheds light on the behavior of nonlinear functions of integrated processes and is thereby useful in the development of an asymptotic theory for regression that involves such nonlinear functions.

Three classes of transformation are explored here: integrable (I) functions, asymptotically homogenous (H) functions and explosive (E) functions. These will be referred to respectively as Classes (I), (H), and (E) in the paper and will be denoted by  $\mathcal{T}(I)$ ,  $\mathcal{T}(H)$  and  $\mathcal{T}(E)$ . More explicitly we define these classes as follows.

**4.1 Definition** A transformation  $T$  is said to be in Class (I), denoted by  $T \in \mathcal{T}(\text{I})$ , iff it is integrable.

**4.2 Definition** A transformation  $T$  is said to be in Class (H), denoted by  $T \in \mathcal{T}(\text{H})$ , iff

$$T(\lambda x) = \nu(\lambda)H(x) + R(x, \lambda)$$

where  $H$  is locally integrable, and  $R$  is such that

(a)  $|R(x, \lambda)| \leq a(\lambda)P(x)$ , where  $\limsup_{\lambda \rightarrow \infty} a(\lambda)/\nu(\lambda) = 0$  and  $P$  is locally integrable, or

(b)  $|R(x, \lambda)| \leq b(\lambda)Q(\lambda x)$ , where  $\limsup_{\lambda \rightarrow \infty} b(\lambda)/\nu(\lambda) < \infty$  and  $Q$  is locally integrable and vanishes at infinity, i.e.,  $Q(x) \rightarrow 0$  as  $|x| \rightarrow \infty$ .

Transformations  $T \in \mathcal{T}(\text{H})$  with  $R$  satisfying conditions (a) and (b) will be said to belong  $\mathcal{T}(\text{H}_1)$  and  $\mathcal{T}(\text{H}_2)$ , respectively.

### 4.3 Remarks

(a) If  $T \in \mathcal{T}(\text{H})$ ,  $T$  has an asymptotically dominating component which is homogenous. All homogenous functions are of this type, and therefore belong to  $\mathcal{T}(\text{H})$  as long as they are locally integrable. If  $T$  is homogenous of degree  $\kappa$ , then we have  $H = T$  and  $\nu(\lambda) = \lambda^\kappa$ . Examples of such functions include  $T(x) = x^\kappa$  for  $\kappa > 0$  and  $T(x) = \text{sgn}(x)$ .

(b) The finite order polynomial given by  $T(x) = x^k + a_1x^{k-1} + \dots + a_k$  for  $k \geq 1$  is in  $\mathcal{T}(\text{H}_1)$  with  $\nu(\lambda) = \lambda^k$  and  $H(x) = x^k$ . For  $a(\lambda) = \lambda^{k-1}|a_1 + a_2/\lambda + \dots + a_k/\lambda^{k-1}|$  and  $P(x) = 1 + |x|^{k-1}$ , we may easily show that  $|R(x, \lambda)| \leq a(\lambda)P(x)$ . Clearly,  $a(\lambda)/\nu(\lambda) \rightarrow 0$  as  $\lambda \rightarrow \infty$ , and  $P$  is locally integrable for  $k \geq 1$ .

(c) The logarithmic function  $T(x) = \log|x|$  belongs to  $\mathcal{T}(\text{H}_1)$ , with the homogenous component given by  $\nu(\lambda) = \log \lambda$  and  $H(x) = 1$ . The residual function then becomes  $R(x, \lambda) = \log|x|$ . To see that it satisfies the above conditions, set  $a(\lambda) = 1$  and  $P(x) = \log|x|$ . Iterated logarithmic functions and polynomials in logarithms are also in  $\mathcal{T}(\text{H})$ , which can be shown similarly.

(d) The distribution function of any random variable belongs to class  $\mathcal{T}(\text{H}_2)$ , with the homogenous component specified by  $\nu(\lambda) = 1$  and  $H(x) = 1\{x \geq 0\}$ . Clearly,  $H$  is locally integrable. If  $T$  is such a function, the residual  $R(x, \lambda)$  is bounded in modulus by  $Q(\lambda x)$ , where  $Q(x) = T(x)1\{x < 0\} + (1 - T(x))1\{x \geq 0\}$ . It is easy to see that  $Q$  is locally integrable and vanishes at infinity. If, in particular, the underlying random variable has finite expectation, then  $Q \in \mathcal{T}(\text{I})$ .

**4.4 Definition** A transformation  $T$  is said to be in Class (E), denoted by  $T \in \mathcal{T}(\text{E})$ , iff

$$T(x) = E(x) + R(x)$$

with  $E$  and  $R$  satisfying the following conditions:

(a)  $E$  is monotone. If  $E$  is increasing (decreasing), then it is positive and differentiable on  $\mathcal{R}_+$  ( $\mathcal{R}_-$ ). Furthermore, if we define  $G(x) = \log E(x)$  on  $\mathcal{R}_+$  ( $\mathcal{R}_-$ )



with derivative  $\dot{G}$ , then as  $\lambda \rightarrow \infty$ ,  $\dot{G}(\lambda x) = \nu(\lambda)D(x) + o(\nu(\lambda))$  uniformly in a neighborhood of  $x$ , where  $D$  is positive (negative) and continuous, and  $\lambda\nu(\lambda) \rightarrow \infty$ .

(b)  $R$  is given such that for any  $x$  and  $y$

$$\frac{\lambda\nu(\lambda)\bar{R}(\lambda x)}{E(\lambda y)} \rightarrow 0$$

as  $\lambda \rightarrow \infty$ , where  $\bar{R}(x) = \sup_{y \leq |x|} |R(y)|$ .

#### 4.5 Remarks

(a) For  $T \in \mathcal{T}(\mathbf{E})$ ,  $E$  denotes the exponential component that is asymptotically dominating. The derivative of the exponent function of  $E$  is assumed to be asymptotically homogenous with base function  $D$  and degree of homogeneity  $\nu$ . If we write  $E(x) = \exp(G(x))$ , then the condition  $\lambda\nu(\lambda) \rightarrow \infty$  ensures that  $G$  increases on  $\mathcal{R}_+$  (or decreases on  $\mathcal{R}_-$ ) faster than the logarithmic function. When there is such an exponential component, all other components with polynomial orders become negligible. They satisfy our conditions for  $R$ , as one may easily check.

(b) The conditions for the exponential component  $E$  of  $T \in \mathcal{T}(\mathbf{E})$  obviously hold for functions like  $E(x) = \exp(x^\kappa)$  for  $\kappa > 0$ , or  $E(x) = x^\kappa e^x \{x > 0\}$  for any finite  $\kappa$ . In the former case, we have  $\nu(\lambda) = \lambda^{\kappa-1}$  and  $D(x) = \kappa x^{\kappa-1}$ . For the latter,  $\nu(\lambda) = 1$  and  $D(x) = 1$ .

**4.6 Theorem** *Let  $T \in \mathcal{T}(\mathbf{I})$ . Then*

$$\frac{1}{\lambda} \int_0^{\lambda^2 t} T(W(r)) dr \xrightarrow{d} \left( \int_{-\infty}^{\infty} T(s) ds \right) L(t, 0)$$

as  $\lambda \rightarrow \infty$ .

**4.7 Theorem** *Let  $T \in \mathcal{T}(\mathbf{H})$  with  $H(\cdot)$  as in Definition 4.2. Then*

$$\frac{1}{\lambda^2 \nu(\lambda)} \int_0^{\lambda^2 t} T(W(r)) dr \xrightarrow{d} \int_{-\infty}^{\infty} H(s) L(t, s) ds$$

as  $\lambda \rightarrow \infty$ .

**4.8 Theorem** *Let  $T \in \mathcal{T}(\mathbf{E})$  with  $\nu$  and  $D$  as in Definition 4.4. Then as  $\lambda \rightarrow \infty$*

$$\frac{\nu(\lambda)}{\lambda T\left(\sup_{0 \leq r \leq \lambda^2 t} W(r)\right)} \int_0^{\lambda^2 t} T(W(r)) dr \xrightarrow{d} \frac{1}{D(s_{\max})} L(t, s_{\max})$$

or

$$\frac{\nu(\lambda)}{\lambda T\left(\inf_{0 \leq r \leq \lambda^2 t} W(r)\right)} \int_0^{\lambda^2 t} T(W(r)) dr \xrightarrow{d} \frac{1}{-D(s_{\min})} L(t, s_{\min})$$

depending upon whether the exponential component  $E$  is increasing or decreasing, and where  $s_{\max} = \sup_{0 \leq r \leq 1} W(r)$  and  $s_{\min} = \inf_{0 \leq r \leq 1} W(r)$ .

#### 4.9 Remarks

(a) Theorems 4.6–4.8 reveal that the asymptotic behavior of the three different types of additive functionals of Brownian motion differ in fundamental ways. For integrable functions, only the local time spent by  $W$  in the vicinity of the origin matters. This is not so for asymptotically homogenous functions, for which the local time of  $W$  at all points contributes to the limit distribution. Finally, the local time that  $W$  spends in the neighborhood of one of its extrema completely determines the asymptotic behavior of an explosive function.

(b) The convergence rates for explosive functions are *path-dependent*, i.e., they depend not only on the size of the sample but also on the actual path of the sample by virtue of the fact that  $\sup_r W(r)$  and  $\inf_r W(r)$  influence the convergence rate.

### 5. Functions of Integrated Processes

Not surprisingly, the moments of functions of integrated processes asymptotically behave rather like the corresponding additive functionals of Brownian motion. We just need some extra conditions to make their limiting behavior invariant with respect to the underlying data generating processes.

**5.1 Theorem** *Suppose  $T \in \mathcal{T}(\mathbb{I})$  and Assumption 2.2 holds with  $p > 4$ . If  $T$  is square integrable and satisfies the Lipschitz condition*

$$|T(x) - T(y)| \leq c|x - y|^\ell$$

*over its support for some constant  $c$  and  $\ell > 6/(p - 2)$ , then*

$$\frac{1}{\sqrt{n}} \sum_{t=1}^n T(x_t) \xrightarrow{d} \left( \int_{-\infty}^{\infty} T(s) ds \right) L(1, 0)$$

*as  $n \rightarrow \infty$ .*

#### 5.2 Remarks

(a) For an indicator function on a bounded set, the result in Theorem 5.1 is applicable as long as  $p > 4$ . The Lipschitz function with  $\ell = 1$  requires, in particular, that  $p > 8$ .

(b) The collection of transformations for which Theorem 5.1 applies is closed under the operation of finite linear combinations. Thus, the result in Theorem 5.1 holds for any piecewise function for which each piece satisfies the given conditions.

**5.3 Theorem** *Let  $T \in \mathcal{T}(\mathbb{H})$  with  $H(\cdot)$  regular. Also, assume that  $T$  is either in  $\mathcal{T}(\mathbb{H}_1)$  with  $P$  locally bounded, or in  $\mathcal{T}(\mathbb{H}_2)$  with  $Q$  bounded and vanishing at infinity. If Assumption 2.1 holds, then*

$$\frac{1}{n\nu(\sqrt{n})} \sum_{t=1}^n T(x_t) \xrightarrow{d} \int_{-\infty}^{\infty} H(s)L(1, s) ds$$

*as  $n \rightarrow \infty$ .*

#### 5.4 Remarks

(a) For Theorem 5.3, we only need Assumption 2.1. This is in contrast to Theorems 5.1 and 5.5 for functions in  $\mathcal{T}(\text{I})$  and  $\mathcal{T}(\text{E})$ , where the stronger Assumption 2.2 is invoked.

(b) The result in Theorem 5.3 is applicable to such functions as  $T(x) = x^\kappa$  for  $\kappa > 0$ ,  $T(x) = \text{sgn}(x)$ ,  $T(x) = x^k + a_1 x^{k-1} + \dots + a_k$  for  $k \geq 1$ , and to all “distribution function”-like transformations.

**5.5 Theorem** *Let  $T \in \mathcal{T}(\text{E})$  and  $\nu(\lambda) = \lambda^m$  with  $m < (p - 8)/6p$ . If Assumption 2.2 holds, then as  $n \rightarrow \infty$*

$$\frac{\nu(\sqrt{n})}{\sqrt{n}T(\max_{1 \leq t \leq n} x_t)} \sum_{t=1}^n T(x_t) \xrightarrow{d} \frac{1}{D(s_{\max})} L(1, s_{\max})$$

or

$$\frac{\nu(\sqrt{n})}{\sqrt{n}T(\min_{1 \leq t \leq n} x_t)} \sum_{t=1}^n T(x_t) \xrightarrow{d} \frac{1}{-D(s_{\min})} L(1, s_{\min})$$

depending upon whether the exponential component  $E$  is increasing or decreasing.

#### 5.6 Remarks

(a) The convergence rates are path-dependent, as in Theorem 4.8, i.e., they depend upon  $\max x_t$  or  $\min x_t$ ,  $t = 1, \dots, n$ , respectively for the increasing and decreasing exponential component of the transformation in  $\mathcal{T}(\text{E})$ .

(b) The result in Theorem 5.5 is applicable for explosive functions such as  $x^\kappa \exp(x)\{x > 0\}$ , as long as  $p > 8$ . However, we only allow functions to be mildly explosive. Functions like  $T(x) = \exp(x^2)$  are excluded. The asymptotic behaviors of such functions may not be invariant, and can be more dependent upon the underlying data generating process.

## 6. Nonlinear Regression Illustrations with Integrated Processes

In this section, we briefly show how to apply the above theory to develop regression asymptotics for models with transformed integrated regressors. Let  $\{x_t\}$  be generated by (1) and (2) and consider the regression model

$$y_t = \alpha f(x_t) + u_t, \tag{10}$$

for  $t = 1, \dots, n$ , where  $\alpha$  is the regression coefficient,  $f$  is a transformation in  $\mathcal{R}$ , and  $\{u_t\}$  are the errors. The least squares estimator  $\hat{\alpha}_n$  of  $\alpha$  in regression (10) is given by

$$\hat{\alpha}_n = \frac{\sum_{t=1}^n f(x_t) y_t}{\sum_{t=1}^n f^2(x_t)} = \alpha + \frac{\sum_{t=1}^n f(x_t) u_t}{\sum_{t=1}^n f^2(x_t)}.$$

When  $f$  is the identity transform, regression (10) reduces to what is known as (a linear) cointegrating regression. Such regressions have become very popular in time

series econometrics following the work of Engle and Granger (1987). However, it is not always clear that the relationship between  $y_t$  and  $x_t$  is linear and such considerations lead naturally to models of the form (10) (just as in the case where  $y_t$  and  $x_t$  are stationary).

Let  $\{\mathcal{F}_t\}$  be the natural filtration for  $\{u_t\}$ , and assume:

- 6.1 Assumption** (a)  $\{u_t\}$  is independent of  $\{w_t\}$ , and  
 (b)  $(u_t, \mathcal{F}_t)$  is a martingale difference sequence with  $\mathbf{E}(u_t^2 | \mathcal{F}_{t-1}) = \sigma^2$  for all  $t$ , and  $\sup_t \mathbf{E}(|u_t|^q | \mathcal{F}_{t-1}) < \infty$  a.s. for some  $q > 2$ .

Assumption 6.1 (a) is stronger than is needed, but is made for simplicity to highlight the effect of the nonlinear transformation on the regression asymptotics. As before, we let  $\sigma^2 = 1$ , since it has only a scaling effect.

The lemma that follows gives the Skorohod embedding of a partial sum and a strong approximation to its quadratic variation as in Phillips and Ploberger (1996). It is useful in the derivation of the regression asymptotics in Theorem 6.3 below.

**6.2 Lemma** *Let Assumption 6.1(b) hold. Then there exists a probability space supporting a standard linear Brownian motion  $U$  and an increasing sequence of stopping times  $\{\tau_t\}_{t \geq 0}$  with  $\tau_0 = 0$  such that  $\frac{1}{\sqrt{n}} \sum_{k=1}^t u_k \stackrel{d}{=} U(\frac{\tau_t}{n})$  and*

$$\sup_{1 \leq t \leq n} \left| \frac{\tau_t - t}{n^\delta} \right| \xrightarrow{a.s.} 0$$

as  $n \rightarrow \infty$  for any  $\delta > \max(1/2, 2/q)$ .

In view of Assumption 6.1(a), we may assume that  $W$  and  $U$  are independent, and defined on a common probability space.

**6.3 Theorem** *Let  $T = f^2$  and denote by  $V$  a standard linear Brownian motion independent of  $W$ . Suppose Assumption 6.1 holds.*

(a) *If  $T$  satisfies the conditions in Theorem 5.1, then as  $n \rightarrow \infty$*

$$\sqrt[4]{n} (\hat{\alpha}_n - \alpha) \xrightarrow{d} \frac{V(1)}{\left( \int_{-\infty}^{\infty} T(s) ds L(1, 0) \right)^{1/2}}$$

(b) *If  $T$  satisfies the conditions in Theorem 5.3, then as  $n \rightarrow \infty$*

$$\sqrt{n\nu(\sqrt{n})} (\hat{\alpha}_n - \alpha) \xrightarrow{d} \frac{V(1)}{\left( \int_{-\infty}^{\infty} H(s)L(1, s) ds \right)^{1/2}}$$

(c) *If  $T$  satisfies the conditions in Theorem 5.5, then as  $n \rightarrow \infty$*

$$\left( \frac{\sqrt{n}T(\max_{1 \leq t \leq n} x_t)}{\nu(\sqrt{n})} \right)^{1/2} (\hat{\alpha}_n - \alpha) \xrightarrow{d} \frac{D(s_{\max})^{1/2}V(1)}{L(1, s_{\max})^{1/2}}$$

or

$$\left( \frac{\sqrt{n}T (\min_{1 \leq t \leq n} x_t)}{\nu(\sqrt{n})} \right)^{1/2} (\hat{\alpha}_n - \alpha) \xrightarrow{d} \frac{-D(s_{\min})^{1/2}V(1)}{L(1, s_{\min})^{1/2}}$$

depending upon whether the exponential component  $E$  is increasing or decreasing.

Theorem 6.3 shows that  $\hat{\alpha}_n$  is consistent when the conditions in Theorems 5.1 and 5.5 are met for  $T = f^2$ . Also, it is consistent if  $T = f^2$  satisfies the conditions in Theorem 5.3 with  $\lambda^2\nu(\lambda) \rightarrow \infty$  as  $\lambda \rightarrow \infty$ . Thus, we may generally expect consistency, in the same way as in other time series regressions under persistent excitation. The limiting distributions are mixed normal, in the same way as for cointegrating regressions (Phillips, 1971). The rate of convergence, however, will vary depending on  $f$ . It can be faster than the convergence rate ( $n$ ) for linear cointegrating regressions, but it can also be slower than the  $\sqrt{n}$  rate for stationary regression. When  $f$  is explosive, as in the case of exponential functions, the convergence rate for  $\hat{\alpha}_n$  is dependent upon the entire sample path of  $x_t$ , as well as the sample size.

Since the sample path of an integrated process typically shows trending behavior, it is interesting to compare (10) with nonlinear regressions on deterministically trending regressors. To be explicit, consider the following two regressions

$$y_t = \frac{\alpha}{|x_t|^\beta} + u_t \quad (11)$$

and

$$y_t = \frac{\alpha}{t^\beta} + u_t \quad (12)$$

where  $\beta > 0$  is a known constant, and the other notation is defined as in (10). The least squares estimators of  $\alpha$  in (11) and (12) are denoted, respectively, by  $\hat{\alpha}_n$  and  $\tilde{\alpha}_n$ . Unlike  $\tilde{\alpha}_n$ ,  $\hat{\alpha}_n$  is not properly defined without some modification, since  $x_t$  may take values in the neighborhood of zero (or could even be zero with positive probability in the case of discrete innovations  $w_t$ ) in which case the regression function is singular. Therefore, we follow the convention introduced in (9), and assume that  $\hat{\alpha}_n$  is computed from a regression on  $x_{nt} = x_t\{|x_t| \geq c_n\} + c_n\{|x_t| < c_n\}$  (in lieu of  $x_t$ ) with  $c_n = n^{-\delta}$  for  $0 < \delta < (p-2)/2p(1+2\beta)$ . See Remark 3.5(b) for our choice of  $c_n$  here. We let Assumption 2.2 hold in the subsequent discussion.

The asymptotic behavior of both  $\hat{\alpha}_n$  and  $\tilde{\alpha}_n$  are critically dependent upon the value of  $\beta$ . For  $0 < \beta < 1/2$ , both  $\hat{\alpha}_n$  and  $\tilde{\alpha}_n$  are consistent, and have limiting distributions given, respectively, by

$$n^{(1-\beta)/2}(\hat{\alpha}_n - \alpha) \xrightarrow{d} \left( \int_0^1 \frac{1}{|W(r)|^{2\beta}} dr \right)^{-1/2} V(1)$$

and

$$n^{1/2-\beta}(\tilde{\alpha}_n - \alpha) \xrightarrow{d} \left( \int_0^1 \frac{1}{r^{2\beta}} dr \right)^{-1/2} V(1).$$

If  $\beta > 1/2$ , however, the asymptotic behavior is very different.

When  $\beta = 1/2$ ,  $(\log n)^{1/2}(\tilde{\alpha}_n - \alpha) \rightarrow_d V(1)$  and  $\tilde{\alpha}_n$  from regression (12) is therefore consistent. The estimator  $\tilde{\alpha}_n$  becomes inconsistent if  $\beta$  exceeds the critical value  $1/2$ , since  $\sum_{t=1}^n 1/t^{2\beta} < \infty$  for  $\beta > 1/2$ , and the excitation condition fails to hold. Faulty intuition here might suggest that regression (12) with  $\beta = 1/2$  is analogous to regression (11) with  $\beta = 1$ , because  $x_t = O_p(\sqrt{t})$ . This might lead to the conjecture that  $\hat{\alpha}_n$  from regression (11) becomes inconsistent when  $\beta > 1$ . Interestingly, however,  $\hat{\alpha}_n$  from regression (11) is consistent for *all* values of  $\beta$ , including  $\beta > 1$ , as shown in the following proposition, which establishes the validity of the excitation condition for the regressor in (11) for all  $\beta$ .

**6.4 Proposition** *Let Assumption 2.2 hold. Then*

$$\sum_{t=1}^n |x_t|^\kappa \xrightarrow{p} \infty$$

as  $n \rightarrow \infty$ , for any  $\kappa \neq -\infty$ .

## 7. Conclusion

The examples given in the previous section involve models that are linear in the parameters and nonlinear in the regressor. Such models are obviously very simple examples of regressions that involve nonlinear functions of integrated processes and our theory therefore provides only a basic extension of cointegrating regression asymptotics. In spite of their simplicity, however, these models do illustrate some important features of more general nonlinear cointegrating regression problems.

First, it is apparent that the signal emanating from a nonstationary regressor can be substantially altered in strength by nonlinear transformations. Moreover, as the strength of the signal is modified, the corresponding rate of convergence of the regression coefficient is affected. Our simple examples show that nonlinear transformations can decrease the rate of convergence over that of a linear cointegrating regression as well as increase this rate. Second, the rate of convergence may in some cases be path dependent, in the sense that the rate itself is stochastic and depends on properties of the process like its maximum or minimum. Finally, the limit theory in all cases considered turns out to be mixed normal, as in linear cointegrating regressions. Indeed, if a Gaussian likelihood approach were adopted, the likelihood would turn out to be in the locally asymptotically mixed normal class, so that an optimal theory of inference can be developed (c.f Jeganathan, 1995, and Phillips, 1991).

Not addressed in this paper is the general task of developing a theory of regression for nonlinear functions of nonstationary regressors in which the parameters also enter in a nonlinear fashion. This task is more complex and of broader scope than what has been completed in this paper, but the results rely intimately on the methods we have introduced here. The results of the broader investigation will be reported by the authors in a subsequent paper.

## 8. Proofs

**8.1 Proof of Lemma 2.3** Parts (a) and (b) are respectively Theorem 3.4 of Phillips and Solo (1992) and Theorem 3 of Akonom (1993).

**8.2 Proof of Lemma 2.4** See e.g., Corollary 7.4 of Chung and Williams (1990).

**8.3 Proof of Lemma 2.5** In what follows, let  $N_n(a, b) = N_n(\nu_n; a, b)$  to simplify notation. For the proof of part (a), we first deduce from Lemma 4 of Akonom (1993) that

$$\mathbf{E} \left( N_n(0, \delta) - \frac{1}{k} N_n(\delta, (k+1)\delta) \right)^2 \leq c \frac{\delta}{n\nu_n} \left( 1 + \frac{k\delta^2 n \log n}{\nu_n^2} \right)$$

and similarly

$$\mathbf{E} \left( N_n(k, (k+1)\delta) - \frac{1}{k} N_n(\delta, (k+1)\delta) \right)^2 \leq c \frac{\delta}{n\nu_n} \left( 1 + \frac{k\delta^2 n \log n}{\nu_n^2} \right)$$

where  $c$  is some constant depending only upon the distribution of  $\{\varepsilon_t\}$  and  $\{\varphi_k\}$ . The stated result now follows immediately since

$$\begin{aligned} & \mathbf{E} (N_n(0, \delta) - N_n(k\delta, (k+1)\delta))^2 \\ & \leq 2 \left( \mathbf{E} \left( N_n(0, \delta) - \frac{1}{k} N_n(\delta, (k+1)\delta) \right)^2 + \mathbf{E} \left( N_n(k, (k+1)\delta) - \frac{1}{k} N_n(\delta, (k+1)\delta) \right)^2 \right) \end{aligned}$$

Part(b) is due to Akonom (1993), Theorem 4.  $\square$

**8.4 Proof of Theorem 3.2** Assume temporarily that  $x_0 = 0$ , and write

$$\frac{1}{n} \sum_{t=1}^n T \left( \frac{x_t}{\sqrt{n}} \right) \stackrel{d}{=} \int_0^1 T(W_n(r)) dr$$

Let  $C = [s_{\min} - 1, s_{\max} + 1]$ , where  $s_{\min}$  and  $s_{\max}$  are defined as in Theorem 4.8. Due to Lemma 2.3 (a), we may take  $n$  sufficiently large so that  $\sup |W_n(r) - W(r)| < \delta_\varepsilon$  for any  $\delta_\varepsilon > 0$ , and that both  $W_n$  and  $W$  are in  $C$  a.s. (Note that  $C$  is path dependent on  $W$  by construction.) Therefore,

$$\underline{T}_\varepsilon(W(r)) \leq T(W_n(r)) \leq \overline{T}_\varepsilon(W(r)) \tag{13}$$

for large  $n$  because of (6). However,

$$\begin{aligned} \int_0^1 (\overline{T}_\varepsilon - \underline{T}_\varepsilon)(W(r)) dr &= \int_{-\infty}^{\infty} (\overline{T}_\varepsilon - \underline{T}_\varepsilon)(s) L(1, s) ds \\ &\leq \left( \sup_s L(1, s) \right) \int_C (\overline{T}_\varepsilon - \underline{T}_\varepsilon)(x) dx \\ &\stackrel{\text{a.s.}}{\rightarrow} 0 \end{aligned} \tag{14}$$

as  $\varepsilon \rightarrow 0$ , due to (7). The stated result now easily follows from (13) and (14). For the case  $x_0 \neq 0$ , simply replace  $W_n$  with  $x_0/\sqrt{n} + W_n$  in the above proof.  $\square$

**8.5 Proof of Theorem 3.4** Again, temporarily assume  $x_0 = 0$ , and write

$$\frac{1}{n} \sum_{t=1}^n T_n \left( \frac{x_t}{\sqrt{n}} \right) \stackrel{d}{=} \int_0^1 T_n(W_n(r)) dr,$$

as in the proof of Theorem 3.2. We define

$$\begin{aligned} A_n &= \left| \int_0^1 T_n(W_n(r)) dr - \int_0^1 T_n(W(r)) dr \right| \\ B_n &= \left| \int_0^1 T_n(W(r)) dr - \int_0^1 T(W(r)) dr \right| \end{aligned}$$

and show

$$\left| \int_0^1 T_n(W_n(r)) dr - \int_0^1 T(W(r)) dr \right| \leq A_n + B_n = o_p(1)$$

below.

Given the conditions on the orders of  $\nu(c_n)$  and  $T(\pm c_n)$ , we may easily deduce from Lemma 2.3(b) and 2.5(b), setting  $\pi_n/\nu_n = c_n$  in the latter, that

$$\begin{aligned} A_n &\leq \nu(c_n) \int_0^1 |W_n(r) - W(r)| dr + |T(\pm c_n)| \\ &\quad \times \left| \int_0^1 1\{|W_n(r)| < c_n\} dr - \int_0^1 1\{|W(r)| < c_n\} dr \right| = o_p(1) \end{aligned} \quad (15)$$

Therefore, it suffices to show that

$$\begin{aligned} B_n &\leq \left| \int_0^1 T(W(r)) 1\{|W(r)| \geq c_n\} - \int_0^1 T(W(r)) dr \right| \\ &\quad + |T(\pm c_n)| \int_0^1 1\{|W(r)| < c_n\} dr = o(1) \text{ a.s.} \end{aligned} \quad (16)$$

It follows from (3) that

$$T(\pm c_n) \int_0^1 1\{|W(r)| \leq c_n\} dr = c_n T(\pm c_n) (L(1, 0) + o(1)) \xrightarrow{\text{a.s.}} 0,$$

since  $T$  is locally integrable and therefore  $c_n T(\pm c_n) \rightarrow 0$  for  $c_n \rightarrow 0$ . Moreover,

$$\begin{aligned} \int_0^1 T(W(r)) 1\{|W(r)| \geq c_n\} dr &= \int_{-\infty}^{\infty} T(s) 1\{|s| \geq c_n\} L(1, s) ds \\ &\stackrel{\text{a.s.}}{\rightarrow} \int_{-\infty}^{\infty} T(s) L(1, s) ds \\ &= \int_0^1 T(W(r)) dr, \end{aligned}$$

by dominated convergence and repeated applications of Lemma 2.4. Notice that  $T(\cdot)1\{|\cdot| \geq c_n\} \rightarrow T(\cdot)$  pointwise except at zero, which is of Lebesgue measure zero. The stated result now follows from (15) and (16).

When  $x_0 \neq 0$ , we may define

$$\begin{aligned} A'_n &= \int_0^1 T_n \left( \frac{x_0}{\sqrt{n}} + W_n(r) \right) dr - \int_0^1 T_n \left( \frac{x_0}{\sqrt{n}} + W(r) \right) dr, \\ B'_n &= \int_0^1 T_n \left( \frac{x_0}{\sqrt{n}} + W(r) \right) dr - \int_0^1 T(W(r)) dr, \end{aligned}$$

instead of  $A_n$  and  $B_n$ , and the stated result holds in the same way.  $\square$



**8.6 Proof of Theorem 4.6** See Proposition 2.2 in Chapter XIII of Revuz and Yor (1994).  $\square$

**8.7 Proof of Theorem 4.7** We have

$$\begin{aligned} \frac{1}{\lambda^2 \nu(\lambda)} \int_0^{\lambda^2 t} T(W(r)) dr &= \frac{1}{\nu(\lambda)} \int_0^t T(W(\lambda^2 r)) dr \\ &\stackrel{d}{=} \frac{1}{\nu(\lambda)} \int_0^t T(\lambda W(r)) dr \\ &= \int_0^t H(W(r)) dr + \frac{1}{\nu(\lambda)} \int_0^t R(W(r), \lambda) dr. \end{aligned}$$

Since  $H$  is assumed to be locally integrable,

$$\int_0^t H(W(r)) dr = \int_{-\infty}^{\infty} H(s) L(t, s) ds,$$

by Lemma 2.4. Therefore, it suffices to show that

$$\frac{1}{\nu(\lambda)} \int_0^t R(W(r), \lambda) dr \xrightarrow{\text{a.s.}} 0$$

to finish the proof.

If  $T \in \mathcal{T}(\mathbb{H}_1)$ , it is immediate that

$$\frac{1}{\nu(\lambda)} \int_0^t |R(W(r), \lambda)| dr \leq \frac{a(\lambda)}{\nu(\lambda)} \int_0^t P(W(r)) dr \xrightarrow{\text{a.s.}} 0,$$

since  $P$  is locally integrable. For  $T \in \mathcal{T}(\mathbb{H}_2)$ , we have from Lemma 2.4

$$\begin{aligned} \frac{1}{\nu(\lambda)} \int_0^t |R(W(r), \lambda)| dr &\leq \frac{b(\lambda)}{\nu(\lambda)} \int_0^t Q(\lambda W(r)) dr \\ &= \frac{b(\lambda)}{\nu(\lambda)} \int_{-\infty}^{\infty} Q(\lambda s) L(t, s) ds. \end{aligned}$$

Since  $Q$  vanishes at infinity,  $Q(\lambda s) \rightarrow 0$  for all  $s$  except  $s = 0$ , which is of Lebesgue measure zero. We may assume w.l.o.g. that  $Q$  is monotone decreasing (increasing) as  $x \rightarrow \infty$  ( $x \rightarrow -\infty$ ), by considering  $Q_*$ ,  $Q_*(x) = \sup_{y \geq |x|} Q(y)$ , in place of  $Q$ , if necessary. Now, for all  $\lambda \geq 1$ ,  $Q(\lambda \cdot)$  is bounded by  $Q(\cdot)$  which is locally integrable. Since  $L(t, \cdot)$  has compact support for any fixed  $t$ , we have

$$\int_{-\infty}^{\infty} Q(\lambda s) L(t, s) ds \xrightarrow{\text{a.s.}} 0, \tag{17}$$

by dominated convergence.  $\square$

**8.8 Proof of Theorem 4.8** We let  $E$  be increasing. The proof for the decreasing  $E$  is quite similar, and omitted. In the proof, we let  $s_{\max} = \bar{s}$  and  $s_{\min} = \underline{s}$  for notational simplicity. Notice first that

$$\begin{aligned} \frac{\nu(\lambda)}{\lambda T \left( \sup_{0 \leq r \leq \lambda^2 t} W(r) \right)} \int_0^{\lambda^2 t} T(W(r)) dr &= \frac{\lambda \nu(\lambda)}{T \left( \sup_{0 \leq r \leq t} W(\lambda^2 r) \right)} \int_0^t T(W(\lambda^2 r)) dr \\ &\stackrel{d}{=} \frac{\lambda \nu(\lambda)}{T(\lambda \bar{s})} \int_0^t T(\lambda W(r)) dr, \end{aligned}$$

for all  $\lambda$ . However, we have

$$\begin{aligned} \frac{\lambda\nu(\lambda)}{T(\lambda\bar{s})} \int_0^t T(\lambda W(r)) dr &= \frac{\lambda\nu(\lambda)}{E(\lambda\bar{s})} \int_0^t T(\lambda W(r)) dr (1 + o(1)) \text{ a.s.} \\ &= \frac{\lambda\nu(\lambda)}{E(\lambda\bar{s})} \int_0^t E(\lambda W(r)) dr (1 + o(1)) \text{ a.s.} \end{aligned} \quad (18)$$

since for  $s_m = \max(\bar{s}, -\underline{s})$

$$\begin{aligned} \frac{|R(\lambda\bar{s})|}{E(\lambda\bar{s})} &\leq \frac{\bar{R}(\lambda s_m)}{E(\lambda\bar{s})} \xrightarrow{\text{a.s.}} 0 \\ \frac{\lambda\nu(\lambda)}{E(\lambda\bar{s})} \int_0^1 |R(\lambda W(r))| dr &\leq \frac{\lambda\nu(\lambda)}{E(\lambda\bar{s})} \bar{R}(\lambda s_m) \xrightarrow{\text{a.s.}} 0, \end{aligned}$$

by the condition on  $R$ .

It follows from Lemma 2.4 that

$$\begin{aligned} \frac{\lambda\nu(\lambda)}{E(\lambda\bar{s})} \int_0^t E(\lambda W(r)) dr &= \frac{\lambda\nu(\lambda)}{E(\lambda\bar{s})} \int_{-\infty}^{\infty} E(\lambda s) L(t, s) ds \\ &= \frac{\lambda\nu(\lambda)}{E(\lambda\bar{s})} \int_0^{\infty} E(\lambda(\bar{s} - s)) L(t, \bar{s} - s) ds. \end{aligned} \quad (19)$$

Now we choose a function  $s(\lambda) \geq 0$  of  $\lambda$  such that

$$s(\lambda) \rightarrow 0 \quad \text{and} \quad \lambda\nu(\lambda)s(\lambda) \rightarrow \infty, \quad (20)$$

as  $\lambda \rightarrow \infty$ . Due to (18) and (19), it suffices to show that

$$\frac{\lambda\nu(\lambda)}{E(\lambda\bar{s})} \int_{s(\lambda)}^{\infty} E(\lambda(\bar{s} - s)) L(1, \bar{s} - s) ds \xrightarrow{\text{a.s.}} 0, \quad (21)$$

$$\frac{\lambda\nu(\lambda)}{E(\lambda\bar{s})} \int_0^{s(\lambda)} E(\lambda(\bar{s} - s)) L(t, \bar{s} - s) ds \xrightarrow{\text{a.s.}} \frac{1}{D(\bar{s})} L(t, \bar{s}), \quad (22)$$

to finish the proof. Note for  $0 \leq s \leq s(\lambda)$  that

$$\begin{aligned} G(\lambda(\bar{s} - s)) - G(\lambda\bar{s}) &= -\lambda s \dot{G}(\lambda(\bar{s} - s_0(\lambda))) \\ &= -\lambda\nu(\lambda)s(D(\bar{s}) + o_{\text{a.s.}}(1)), \end{aligned} \quad (23)$$

uniformly in  $s$  for large  $\lambda$ , where  $0 \leq s_0(\lambda) \leq s(\lambda)$ . By (20),  $s(\lambda), s_0(\lambda) \rightarrow 0$  as  $\lambda \rightarrow \infty$ .

Subsequently using the fact that  $E$  is increasing and  $\int_{-\infty}^{\infty} L(t, s) ds = t$ , along with (23), we have

$$\begin{aligned} &\frac{\lambda\nu(\lambda)}{E(\lambda\bar{s})} \int_{s(\lambda)}^{\infty} E(\lambda(\bar{s} - s)) L(t, \bar{s} - s) ds \\ &\leq \lambda\nu(\lambda)t \frac{E(\lambda(\bar{s} - s(\lambda)))}{E(\lambda\bar{s})} \\ &= \lambda\nu(\lambda)t \exp(G(\lambda(\bar{s} - s(\lambda))) - G(\lambda\bar{s})) \\ &= \lambda\nu(\lambda)t \exp(-\lambda\nu(\lambda)s(\lambda)(D(\bar{s}) + o(1))) \xrightarrow{\text{a.s.}} 0, \end{aligned}$$

as  $\lambda \rightarrow \infty$ , since  $\lambda\nu(\lambda)s(\lambda) \rightarrow \infty$  by (20) and  $D(\bar{s}) > 0$ . This shows (21). Now, by (23) again,

$$\begin{aligned}
& \frac{\lambda\nu(\lambda)}{E(\lambda\bar{s})} \int_0^{s(\lambda)} E(\lambda(\bar{s} - s)) L(t, \bar{s} - s) ds \\
&= \lambda\nu(\lambda) \int_0^{s(\lambda)} \exp(-\lambda\nu(\lambda)sD(\bar{s})(1 + o(1))) L(t, \bar{s} - s) ds \\
&= L(t, \bar{s}) \lambda\nu(\lambda) \int_0^{s(\lambda)} \exp(-\lambda\nu(\lambda)sD(\bar{s})) ds (1 + o(1)) \\
&= L(t, \bar{s}) \int_0^{\lambda\nu(\lambda)s(\lambda)} \exp(-sD(\bar{s})) ds (1 + o(1)) \\
&\xrightarrow{\text{a.s.}} \frac{1}{D(\bar{s})} L(t, \bar{s}),
\end{aligned}$$

and this proves (22).  $\square$

**8.9 Proof of Theorem 5.1** Assume  $x_0 = 0$  and write

$$\frac{1}{\sqrt{n}} \sum_{t=1}^n T(x_t) =_d \sqrt{n} \int_0^1 T(\sqrt{n}W_n(r)) dr$$

If  $x_0 \neq 0$ , then we may consider the function  $T(\cdot + x_0)$  in place of  $T(\cdot)$ . It is easy to see that all the proofs go through under this replacement.

Now let

$$\kappa_n = n^a \quad \text{and} \quad \delta_n = n^{-b} \tag{24}$$

for  $a, b > 0$  satisfying

$$a - (1 + \ell)b < 0 \tag{25}$$

$$(6b - 1)p + 2 < 0 \tag{26}$$

$$2a - 1 < 0 \tag{27}$$

$$4a - 4b - 1 < 0 \tag{28}$$

$$(a - b)p - 1 > 0 \tag{29}$$

and define  $T_n$ ,  $T'_n$  and  $T''_n$  by

$$T_n(x) = T(x)1\{-\kappa_n\delta_n \leq x < \kappa_n\delta_n\}$$

$$T'_n(x) = T(x)1\{x \geq \kappa_n\delta_n\}$$

$$T''_n(x) = T(x)1\{x < -\kappa_n\delta_n\}$$

so that  $T = T_n + T'_n + T''_n$ . We will show that

$$\sqrt{n} \int_0^1 T_n(\sqrt{n}W_n(r)) dr = \left( \int_{-\infty}^{\infty} T(s) ds \right) L(1, 0) + o_p(1), \tag{30}$$

and

$$\sqrt{n} \int_0^1 T'_n(\sqrt{n}W_n(r)) dr = o_p(1), \tag{31}$$

$$\sqrt{n} \int_0^1 T''_n(\sqrt{n}W_n(r)) dr = o_p(1), \tag{32}$$

from which the stated result follows directly. For notational brevity, set  $\nu_n = \sqrt{n}$ , and let  $N_n(a, b) = N_n(\nu_n; a, b)$  and  $N(a, b) = N(\nu_n; a, b)$  in what follows, for  $N_n$  and  $N$  defined in (4) and (5).

To show (30), we first define

$$T_{\delta_n}(x) = \sum_{k=-\kappa_n}^{\kappa_n-1} T(k\delta_n) 1_{\{k\delta_n \leq x < (k+1)\delta_n\}}.$$

It follows from the Lipschitz condition for  $T$  that  $\sup |T_n(x) - T_{\delta_n}(x)| \leq \delta_n^\ell$ , and therefore,

$$\begin{aligned} & \left| \sqrt{n} \int_0^1 T_n(\sqrt{n}W_n(r)) dr - \sqrt{n} \int_0^1 T_{\delta_n}(\sqrt{n}W_n(r)) dr \right| \\ & \leq \kappa_n \delta_n^{1+\ell} \left( \frac{\sqrt{n}}{\kappa_n \delta_n} N_n(-\kappa_n \delta_n, \kappa_n \delta_n) \right) = O_p(\kappa_n \delta_n^{1+\ell}) = o_p(1), \end{aligned} \quad (33)$$

given the conditions for  $\kappa_n$  and  $\delta_n$  in (24) and (25). Note that

$$\frac{\sqrt{n}}{\kappa_n \delta_n} N_n(-\kappa_n \delta_n, \kappa_n \delta_n) = 2L(1, 0) + o_p(1),$$

under condition (26), due to Lemma 2.5(b).

Now,

$$\begin{aligned} \sqrt{n} \int_0^1 T_{\delta_n}(\sqrt{n}W_n(r)) dr &= \sqrt{n} \sum_{k=-\kappa_n}^{\kappa_n-1} T(k\delta_n) N_n(k\delta_n, (k+1)\delta_n) \\ &= \sqrt{n} \left( \sum_{k=-\kappa_n}^{\kappa_n-1} T(k\delta_n) \right) N_n(0, \delta_n) + R_n, \end{aligned} \quad (34)$$

where

$$R_n = \sqrt{n} \sum_{k=-\kappa_n}^{\kappa_n-1} T(k\delta_n) (N_n(k\delta_n, (k+1)\delta_n) - N_n(0, \delta_n)).$$

It follows from the Cauchy-Schwarz inequality and Lemma 2.5(a) that

$$\begin{aligned} \mathbf{E}(R_n^2) &\leq n \left( \sum_{k=-\kappa_n}^{\kappa_n-1} T(k\delta_n)^2 \right) \sum_{k=-\kappa_n}^{\kappa_n-1} \mathbf{E}(N_n(0, \delta_n) - N_n(k\delta_n, (k+1)\delta_n))^2 \\ &\leq \left( \int_{-\infty}^{\infty} T_{\delta_n}^2(s) ds \right) \left( c_1 \frac{\kappa_n}{\sqrt{n}} + c_2 \frac{\kappa_n^2 \delta_n^2 \log n}{\sqrt{n}} \right) = o(1), \end{aligned}$$

due to the conditions for  $\kappa_n$  and  $\delta_n$  in (24), (27) and (28), and where  $c_1$  and  $c_2$  are some constants.

However, we have

$$\begin{aligned} \sqrt{n} \left( \sum_{k=-\kappa_n}^{\kappa_n-1} T(k\delta_n) \right) N_n(0, \delta_n) &= \left( \int_{-\infty}^{\infty} T_{\delta_n}(s) ds \right) \frac{\sqrt{n}}{\delta_n} N_n(0, \delta_n) \\ &= \left( \int_{-\infty}^{\infty} T(s) ds \right) L(1, 0) + o_p(1), \end{aligned} \quad (35)$$

due to (26) for  $\kappa_n$  and  $\delta_n$  in (24). Notice that

$$\frac{\sqrt{n}}{\delta_n} N_n(0, \delta_n) \xrightarrow{p} L(1, 0),$$

under condition (26), by Lemma 2.5 (b). Moreover,

$$\begin{aligned} \int_{-\infty}^{\infty} T_{\delta_n}(s) ds &= \int_{-\infty}^{\infty} T_n(s) ds + o\left(\kappa_n \delta_n^2\right) \\ \int_{-\infty}^{\infty} T_n(s) ds &= \int_{-\infty}^{\infty} T(s) ds + o(1), \end{aligned}$$

We now have (30) from (33), (34) and (35).

Next we show (31) and (32). Let

$$\varepsilon_n = \sup_{0 \leq r \leq 1} |W_n(r) - W(r)|. \quad (36)$$

By taking  $n$  sufficiently large, we may assume that  $T'_n$  and  $T''_n$  are monotone (decreasing and increasing, respectively) on their supports. This causes no loss in generality, since we may always bound  $T'_n$  and  $T''_n$  by such functions if  $T$  is integrable. Therefore,

$$\begin{aligned} T'_n(\sqrt{n}W_n(r)) &\leq T(\sqrt{n}(W(r) - \varepsilon_n)) 1_{\{\sqrt{n}(W_n(r) + \varepsilon_n) > \kappa_n \delta_n\}} \\ T''_n(\sqrt{n}W_n(r)) &\leq T(\sqrt{n}(W(r) + \varepsilon_n)) 1_{\{\sqrt{n}(W_n(r) - \varepsilon_n) < -\kappa_n \delta_n\}} \end{aligned}$$

It follows that

$$\begin{aligned} &\sqrt{n} \int_0^1 T'_n(\sqrt{n}W_n(r)) dr \\ &\leq \sqrt{n} \int_0^1 T(\sqrt{n}(W(r) - \varepsilon_n)) 1_{\{\sqrt{n}(W(r) + \varepsilon_n) > \kappa_n \delta_n\}} dr \\ &= \sqrt{n} \int_{-\infty}^{\infty} (T\sqrt{n}(s - \varepsilon_n)) 1_{\{\sqrt{n}(s + \varepsilon_n) > \kappa_n \delta_n\}} L(1, s) ds \\ &= \int_{-\infty}^{\infty} T(s) 1_{\{s > \kappa_n \delta_n - 2\sqrt{n}\varepsilon_n\}} L\left(1, \frac{s}{\sqrt{n}} + \varepsilon_n\right) ds \xrightarrow{p} 0, \end{aligned}$$

since  $\kappa_n \delta_n - 2\sqrt{n}\varepsilon_n \rightarrow_p \infty$ , due to (24) and (29). Similarly,

$$\begin{aligned} &\sqrt{n} \int_0^1 T''_n(\sqrt{n}W_n(r)) dr \\ &\leq \sqrt{n} \int_0^1 T(\sqrt{n}(W(r) + \varepsilon_n)) 1_{\{\sqrt{n}(W(r) - \varepsilon_n) > \kappa_n \delta_n\}} dr \\ &= \sqrt{n} \int_{-\infty}^{\infty} T(\sqrt{n}(s + \varepsilon_n)) 1_{\{\sqrt{n}(s - \varepsilon_n) > \kappa_n \delta_n\}} L(1, s) ds \\ &= \int_{-\infty}^{\infty} T(s) 1_{\{s < -\kappa_n \delta_n + 2\sqrt{n}\varepsilon_n\}} L\left(1, \frac{s}{\sqrt{n}} - \varepsilon_n\right) ds \xrightarrow{p} 0, \end{aligned}$$

since  $-\kappa_n \delta_n + 2\sqrt{n}\varepsilon_n \rightarrow_p -\infty$ , again due to (24) and (29). The proof is therefore complete.  $\square$

**8.10 Proof of Theorem 5.3** Again let  $x_0 = 0$  for simplicity. The proof for  $x_0 \neq 0$  is the same with only  $W_n(r)$  being replaced by  $x_0/\sqrt{n} + W_n(r)$  in what follows. Write

$$\begin{aligned} \frac{1}{n\nu(\sqrt{n})} \sum_{t=1}^n T(x_t) &\stackrel{d}{=} \frac{1}{\nu(\sqrt{n})} \int_0^1 T(\sqrt{n}W_n(r)) dr \\ &= \int_0^1 H(W_n(r)) dr + \frac{1}{\nu(\sqrt{n})} \int_0^1 R(W_n(r), \sqrt{n}) dr. \end{aligned}$$

Since  $H$  is regular, it follows that

$$\int_0^1 H(W_n(r)) dr \xrightarrow{\text{a.s.}} \int_0^1 H(W(r)) dr = \int_{-\infty}^{\infty} H(s)L(1, s)ds.$$

It therefore suffices to show

$$\frac{1}{\nu(\sqrt{n})} \int_0^1 R(W_n(r), \sqrt{n}) dr \xrightarrow{\text{a.s.}} 0$$

to complete the proof.

If  $T \in \mathcal{T}(H_1)$ , it follows immediately that

$$\frac{1}{\nu(\sqrt{n})} \int_0^1 |R(W_n(r), \sqrt{n})| dr \leq \frac{a(\sqrt{n})}{\nu(\sqrt{n})} \int_0^1 P(W_n(r)) dr \xrightarrow{\text{a.s.}} 0,$$

since  $P$  is locally bounded. For  $T \in \mathcal{T}(H_2)$ , we need to show

$$\frac{1}{\nu(\sqrt{n})} \int_0^1 |R(W_n(r), \sqrt{n})| dr \leq \frac{b(\sqrt{n})}{\nu(\sqrt{n})} \int_0^1 Q(\sqrt{n}W_n(r)) dr \xrightarrow{\text{a.s.}} 0, \quad (37)$$

where  $Q$  is bounded and vanishes at infinity. We may assume w.l.o.g. that  $Q$  is monotone decreasing (increasing) for  $x > 0$  ( $x < 0$ ), as noted in the proof of Theorem 4.7. We may thus write  $Q = Q_1 - Q_2$  with both  $Q_1$  and  $Q_2$  bounded and nondecreasing, and let  $\varepsilon_n$  be defined as in (36). It follows that

$$\begin{aligned} & Q_1(\sqrt{n}(W(r) - \varepsilon_n)) - Q_2(\sqrt{n}(W(r) + \varepsilon_n)) \\ & \leq Q(\sqrt{n}W_n(r)) \\ & \leq Q_1(\sqrt{n}(W(r) + \varepsilon_n)) - Q_2(\sqrt{n}(W(r) - \varepsilon_n)). \end{aligned} \quad (38)$$

However,

$$\begin{aligned} \int_0^1 Q_i(\sqrt{n}(W(r) \pm \varepsilon_n)) dr &= \int_{-\infty}^{\infty} Q_i(\sqrt{n}(s \pm \varepsilon_n)) L(1, s) ds \\ &= \int_{-\infty}^{\infty} Q_i(\sqrt{n}s) L(1, s \mp \varepsilon_n) ds \\ &= \int_{-\infty}^{\infty} Q_i(\sqrt{n}s) L(1, s) ds (1 + o(1)) \text{ a.s. } , \end{aligned}$$

since the  $Q_i$ 's are bounded and  $L(1, \cdot)$  is continuous. Therefore,

$$\begin{aligned} & \int_0^1 (Q_1(\sqrt{n}(W(r) \mp \varepsilon_n)) - Q_2(\sqrt{n}(W(r) \pm \varepsilon_n))) dr \\ &= \int_{-\infty}^{\infty} (Q_1(\sqrt{n}(s \mp \varepsilon_n)) - Q_2(\sqrt{n}(s \pm \varepsilon_n))) L(1, s) ds \\ &= \int_{-\infty}^{\infty} Q(\sqrt{n}s) L(1, s) ds (1 + o(1)) \text{ a.s. } . \end{aligned} \quad (39)$$

Now (37) follows easily from (38) and (39), due to (17).  $\square$

**8.11 Proof of Theorem 5.5** Let  $E$  be increasing, and let  $\bar{s}_n = \sup W_n(r)$  and  $\bar{s} = \sup W(r)$ . For simplicity, assume  $x_0 = 0$ . For the case  $x_0 \neq 0$ , we replace  $W_n(r)$  and  $\bar{s}_n$  respectively by  $W_n(r) + x_0/\sqrt{n}$  and  $\bar{s}_n + x_0/\sqrt{n}$  in what follows. All the proofs go through with this replacement. Write

$$\frac{\nu(\sqrt{n})}{\sqrt{n}T(\max_{1 \leq t \leq n} x_t)} \sum_{t=1}^n T(x_t) \stackrel{d}{=} \frac{\sqrt{n}\nu(\sqrt{n})}{T(\sqrt{n}\bar{s}_n)} \int_0^1 T(\sqrt{n}W_n(r)) dr,$$

and notice that

$$\frac{\sqrt{n}\nu(\sqrt{n})}{T(\sqrt{n}\bar{s}_n)} \int_0^1 T(\sqrt{n}W_n(r)) dr = \frac{\sqrt{n}\nu(\sqrt{n})}{E(\sqrt{n}\bar{s}_n)} \int_0^1 E(\sqrt{n}W_n(r)) dr (1 + o_p(1)), \quad (40)$$

which we can show in the same way as (18) in the proof of Theorem 4.8.

Let  $\nu_n = \sqrt{n}\nu(\sqrt{n})$ , and let  $s_n$  be a sequence of numbers such that  $s_n \rightarrow 0$  and  $\nu_n s_n \rightarrow \infty$ . Since  $\bar{s}_n \rightarrow_p \bar{s}$  and  $s_n \rightarrow 0$ , we have similar to (23) in the proof of Theorem 4.8

$$G(\sqrt{n}(\bar{s}_n - s)) - G(\sqrt{n}\bar{s}_n) = -\nu_n s (D(\bar{s}) + o_p(1)), \quad (41)$$

uniformly in  $s \in [0, s_n]$ , for sufficiently large  $n$ . Therefore, if we write

$$\frac{\nu_n}{E(\sqrt{n}\bar{s}_n)} \int_0^1 E(\sqrt{n}W_n(r)) dr = A_n + B_n, \quad (42)$$

where

$$\begin{aligned} A_n &= \frac{\nu_n}{E(\sqrt{n}\bar{s}_n)} \int_0^1 E(\sqrt{n}W_n(r)) \{W_n(r) \geq \bar{s}_n - s_n\} dr, \\ B_n &= \frac{\nu_n}{E(\sqrt{n}\bar{s}_n)} \int_0^1 E(\sqrt{n}W_n(r)) \{W_n(r) < \bar{s}_n - s_n\} dr, \end{aligned}$$

then it follows from (41) that

$$\begin{aligned} A_n &= \nu_n \int_0^1 \exp(-\nu_n D(\bar{s})(\bar{s}_n - W_n(r))) dr (1 + o_p(1)), \\ B_n &= o_p(1), \end{aligned}$$

in parallel to (21) and (22) in the proof of Theorem 4.8.

Define  $W'_n$  and  $W'$  by

$$W'_n(r) = \bar{s}_n - W_n(r) \quad \text{and} \quad W'(r) = \bar{s} - W(r),$$

i.e., Brownian motion reflected at the supremum and its sample analogue. Denote by  $L'$  the local time of  $W'$ . Furthermore, we define  $N'_n$  and  $N'$  for  $W'_n$  and  $W'$  in the same way as  $N_n$  and  $N$  for  $W_n$  and  $W$  given in (4) and (5), respectively. Write  $N'_n(a, b) = N'_n(\nu_n; a, b)$  and  $N'(a, b) = N'(\nu_n; a, b)$  for short. Though we do not provide the details, it is obvious that all the results in Akonom (1993), and therefore our Lemmas 2.3 and 2.5 hold for  $W'_n$  and  $W'$ , as well as  $W_n$  and  $W$ .

Now we write

$$A_n = \nu_n \int_0^1 F(\nu_n W'_n(r)) dr (1 + o_p(1)),$$

with

$$F(x) = e^{-xD(\bar{s})} \{x \geq 0\}.$$

To analyze  $A_n$ , we define  $\kappa_n$  and  $\delta_n$  as in (24) with  $a$  and  $b$  satisfying

$$a - 2b < 0 \quad (43)$$

$$2a + m - 1 < 0 \quad (44)$$

$$4a - 4b - m - 1 < 0 \quad (45)$$

$$(6b + 3m - 1)p + 2 < 0 \quad (46)$$

$$(2a - 2b - m)p - 2 > 0 \quad (47)$$

and let  $\nu_n s_n = \kappa_n \delta_n$ . It is tedious but straightforward to check that  $a$  and  $b$  satisfying all (43)–(47) exist, given our conditions on  $m$  and  $p$ .

We decompose  $F$  into  $F_n$  and  $F'_n$ , where

$$\begin{aligned} F_n(x) &= e^{-x D(\bar{s})} \mathbf{1}\{0 \leq x < \kappa_n \delta_n\}, \\ F'_n(x) &= e^{-x D(\bar{s})} \mathbf{1}\{x \geq \kappa_n \delta_n\}. \end{aligned}$$

It will be shown that

$$\nu_n \int_0^1 F_n(\nu_n W'_n(r)) dr = \left( \int_{-\infty}^{\infty} F(s) ds \right) L'(0, 1) + o_p(1), \quad (48)$$

$$\nu_n \int_0^1 F'_n(\nu_n W'_n(r)) dr = o_p(1), \quad (49)$$

from which we may easily deduce the stated result, upon noticing that

$$\int_{-\infty}^{\infty} F(s) ds = \frac{1}{D(\bar{s})} \quad \text{and} \quad L'(1, 0) = L(1, \bar{s}),$$

together with (40) and (42).

To show (48), we first introduce

$$F_{\delta_n}(x) = \sum_{k=0}^{\kappa_n-1} e^{-k \delta_n D(\bar{s})} \mathbf{1}\{k \delta_n \leq x < (k+1) \delta_n\},$$

and notice that

$$\begin{aligned} \left| \nu_n \int_0^1 F_n(\nu_n W'_n(r)) dr - \nu_n \int_0^1 F_{\delta_n}(\nu_n W'_n(r)) dr \right| &\leq \kappa_n \delta_n^2 \left( \frac{\nu_n}{\kappa_n \delta_n} N'_n(0, \kappa_n \delta_n) \right) \\ &= O_p(\kappa_n \delta_n^2) = o_p(1), \end{aligned} \quad (50)$$

under conditions (43) and (46). Note that

$$\frac{\nu_n}{\kappa_n \delta_n} N'_n(0, \kappa_n \delta_n) = L'(0, 1) + o_p(1),$$

under condition (46) by Lemma 2.5(b).

Secondly,

$$\begin{aligned} \nu_n \int_0^1 F_{\delta_n}(\nu_n W'_n(r)) dr &= \nu_n \sum_{k=0}^{\kappa_n-1} e^{-k \delta_n D(\bar{s})} N'_n(k \delta_n, (k+1) \delta_n) \\ &= \nu_n \left( \sum_{k=0}^{\kappa_n-1} e^{-k \delta_n D(\bar{s})} \right) N'_n(0, \delta_n) + R'_n, \end{aligned} \quad (51)$$

where

$$R'_n = \nu_n \sum_{k=0}^{\kappa_n-1} e^{-k \delta_n D(\bar{s})} (N'_n(k \delta_n, (k+1) \delta_n) - N'_n(0, \delta_n)),$$



and therefore,

$$\begin{aligned} \mathbf{E}(R_n'^2) &\leq \nu_n^2 \left( \sum_{k=0}^{\kappa_n-1} e^{-2k\delta_n D(\bar{s})} \right) \sum_{k=0}^{\kappa_n-1} \mathbf{E} \left( N_n'(k\delta_n, (k+1)\delta_n) - N_n'(0, \delta_n) \right)^2 \\ &\leq \left( \int_{-\infty}^{\infty} F_{\delta_n}^2(s) ds \right) \left( c_1 \frac{\nu_n \kappa_n}{n} + c_2 \frac{\delta_n^2 \kappa_n^2 \log n}{\nu_n} \right) \rightarrow 0, \end{aligned}$$

by conditions (44) and (45), where  $c_1$  and  $c_2$  are some constants.

Thirdly,

$$\begin{aligned} \nu_n \left( \sum_{k=0}^{\kappa_n-1} e^{-k\delta_n D(\bar{s})} \right) N_n'(0, \delta_n) &= \left( \int_{-\infty}^{\infty} F_{\delta_n}(s) ds \right) \left( \frac{\nu_n}{\delta_n} N_n'(0, \delta_n) \right) \\ &= \left( \int_{-\infty}^{\infty} F(s) ds \right) L'(1, 0) + o_p(1). \end{aligned} \quad (52)$$

Notice that

$$\begin{aligned} \int_{-\infty}^{\infty} F_{\delta_n}(s) ds &= \int_{-\infty}^{\infty} F_n(s) ds + O(\kappa_n \delta_n^2), \\ \int_{-\infty}^{\infty} F_n(s) ds &= \int_{-\infty}^{\infty} F(s) ds + O(e^{-\kappa_n \delta_n}). \end{aligned}$$

Also, by Lemma 2.5(b)

$$\frac{\nu_n}{\delta_n} N_n'(0, \delta_n) = L'(1, 0) + o_p(1),$$

under condition (46). Then (48) follows from (50), (51) and (52).

Finally, for  $\varepsilon_n$  defined in (36)

$$\begin{aligned} \nu_n \int_0^1 F_n'(\nu_n W_n'(r)) dr &\leq \nu_n \int_0^1 F(\nu_n(W'(r) - \varepsilon_n)) \mathbf{1}_{\{\nu_n(W'(r) + \varepsilon_n) > \kappa_n \delta_n\}} dr \\ &= \nu_n \int_{-\infty}^{\infty} F(\nu_n(s - \varepsilon_n)) \mathbf{1}_{\{\nu_n(s + \varepsilon_n) > \kappa_n \delta_n\}} L'(1, s) ds \\ &= \int_{-\infty}^{\infty} F(s) \mathbf{1}_{\{s > \kappa_n \delta_n - \nu_n \varepsilon_n\}} L' \left( 1, \frac{s}{\nu_n} + \varepsilon_n \right) ds \xrightarrow{p} 0, \end{aligned}$$

since  $\kappa_n \delta_n - \nu_n \varepsilon_n \xrightarrow{p} \infty$  under condition (47), which proves (49). The proof is therefore complete.  $\square$

**8.12 Proof of Lemma 6.2** By Theorem A1, page 269 of Hall and Heyde (1980) there exist a probability space  $(\Omega, \mathbf{P}, \mathcal{F})$  supporting  $\{U_t\}$ ,  $U_t = \sum_{k=1}^t u_k$ , a Brownian motion  $U$  with variance  $\sigma^2$  and a time change  $\{\tau_t\}$  such that

- (a)  $\tau_t$  is  $\mathcal{F}_t$ -measurable,
- (b)  $\mathbf{E}((\Delta\tau_t)^r | \mathcal{F}_{t-1}) \leq \mathbf{E}(|u_t|^{2r} | \mathcal{F}_{t-1})$  a.s. for  $r \geq 1$ , and
- (c)  $\mathbf{E}(\Delta\tau_t | \mathcal{F}_{t-1}) = 1$ ,

where  $\mathcal{F}_t$  is the  $\sigma$ -field generated by  $(U_k)_{k=1}^t$  and  $U(r)$  for  $0 \leq r \leq \tau_t$ .

Let  $1 \leq r \leq \min(2, q/2)$ . Then we have

$$\mathbf{E}(|\Delta\tau_t - 1|^r | \mathcal{F}_{t-1}) \leq c \sup_{t \geq 1} \mathbf{E}(|u_t|^q | \mathcal{F}_{t-1}) < \infty \quad \text{a.s.}$$

for some constant  $c$ . Therefore,

$$\sum_{t=1}^{\infty} t^{-r\delta} \mathbf{E}(|\Delta\tau_t - 1|^r | \mathcal{F}_{t-1}) < \infty \quad \text{a.s.}$$

since  $r\delta > 1$ , and we have from Theorem 2.18 of Hall and Heyde (1980) that

$$\frac{\tau_t - t}{t^\delta} \xrightarrow{\text{a.s.}} 0,$$

as  $t \rightarrow \infty$  for  $\delta > \max(1/2, 2/q)$ . Therefore, for any  $\varepsilon > 0$  given, there exists  $n'$  such that  $|\tau_t - t|/t^\delta < \varepsilon$  for all  $t > n'$ . Choose  $n \geq n'$  such that  $n > (\max_{1 \leq t \leq n'} |\tau_t - t|/\varepsilon)^{1/\delta}$ . It is easy to check

$$\sup_{1 \leq t \leq n} \left| \frac{\tau_t - t}{n^\delta} \right| < \varepsilon \quad \text{a.s.}$$

as was to be shown.  $\square$

**8.13 Proof of Theorem 6.3** To prove part (a), construct the process

$$\begin{aligned} M_n(r) &= \sqrt[4]{n} \sum_{t=1}^{k-1} f(\sqrt{n}W_n(\frac{t}{n})) (U(\frac{\tau_t}{n}) - U(\frac{\tau_{t-1}}{n})) \\ &\quad + \sqrt[4]{n} f(\sqrt{n}W_n(\frac{k}{n})) (U(r) - U(\frac{\tau_{k-1}}{n})), \end{aligned} \quad (53)$$

for  $\tau_{k-1}/n < r \leq \tau_k/n$ ,  $k = 1, \dots, n$ . Note that  $M_n$  is a continuous martingale such that

$$\frac{1}{\sqrt[4]{n}} \sum_{t=1}^n f(x_t) u_t \stackrel{d}{=} M_n\left(\frac{\tau_n}{n}\right).$$

The quadratic variation process  $[M_n]$  of  $M_n$  is given by

$$\begin{aligned} [M_n]_r &= \sqrt{n} \sum_{t=1}^{k-1} f^2(\sqrt{n}W_n(\frac{t}{n})) (\frac{\tau_t}{n} - \frac{\tau_{t-1}}{n}) \\ &\quad + \sqrt{n} f^2(\sqrt{n}W_n(\frac{k}{n})) (r - \frac{\tau_{k-1}}{n}) \\ &= \sqrt{n} \int_0^r f^2(\sqrt{n}W_n(s)) ds + o_p(1), \end{aligned}$$

since

$$\sup_{1 \leq t \leq n} \left| \left( \frac{\tau_t}{n} - \frac{\tau_{t-1}}{n} \right) - \frac{1}{n} \right| = o(1) \quad \text{a.s.}$$

due to Lemma 6.2. Therefore,

$$[M_n]_r \xrightarrow{p} \left( \int_{-\infty}^{\infty} T(s) ds \right) L(r, 0), \quad (54)$$

as shown in the proof of Theorem 5.1. Moreover, if we denote by  $[M_n, W]$  the covariation process of  $M_n$  and  $W$ , then

$$[M_n, W]_r = 0 \quad (55)$$

for all  $r \in [0, 1]$ , due to the independence of  $U$  and  $W$ . The asymptotic distribution of the continuous martingale  $M_n$  in (53) is completely determined by (54) and (55), as shown in Revuz and Yor (1994, Theorem 2.3, p. 496).

Now define the sequence of time changes

$$\rho_n(r) = \inf \{s \mid [M_n]_s > r\}$$

and subsequently set

$$V_n(r) = M_n(\rho_n(r)).$$

The process  $V_n$  is the DDS (or Dambis, Dubins–Schwarz) Brownian motion of the continuous martingale  $M_n$  [see, for example, Revuz and Yor (1994), Theorem 1.6, p. 173]. It follows that  $(V_n, W)$  converges jointly in distribution to two independent standard linear Brownian motions  $(V, W)$ , say. Therefore,

$$\begin{aligned} M_n\left(\frac{\tau_n}{n}\right) &= M_n(1) + o_p(1) \\ &\xrightarrow{d} V\left(\int_{-\infty}^{\infty} T(s) ds L(1, 0)\right), \end{aligned}$$

which gives the result stated in (a). The proofs for (b) and (c) are similar, and are therefore omitted.  $\square$

**8.14 Proof of Proposition 6.4** The case  $\kappa \geq 0$  is straightforward because

$$n^{-1-\kappa/2} \sum_{t=1}^n |x_t|^\kappa = \frac{1}{n} \sum_{t=1}^n \left| \frac{x_t}{\sqrt{n}} \right|^\kappa \xrightarrow{d} \int_0^1 |W(r)|^\kappa dr, \quad (56)$$

by Theorem 3.2, since  $T(x) = |x|^\kappa$  is regular. In the case where  $-1 < \kappa < 0$  Theorem 3.4 is applicable and (56) again yields the stated result. For the case  $\kappa \leq -1$  we use a different argument. Bound  $\sum_{t=1}^n |x_t|^\kappa$  below as

$$\begin{aligned} \sum_{t=1}^n |x_t|^\kappa &\geq \sum_{t=1}^n |x_t|^\kappa \mathbf{1} \left\{ \frac{|x_t - x_0|}{\sqrt{n}} \leq \frac{1}{\sqrt{n}} \right\} \\ &\geq (1 + |x_0|)^\kappa \sum_{t=1}^n \mathbf{1} \left\{ \frac{|x_t - x_0|}{\sqrt{n}} \leq \frac{1}{\sqrt{n}} \right\} \\ &\stackrel{d}{=} n (1 + |x_0|)^\kappa \int_0^1 \mathbf{1} \left\{ |W_n(r)| \leq \frac{1}{\sqrt{n}} \right\} dr. \end{aligned}$$

Then, from Lemma 2.5(b)

$$\begin{aligned} \sqrt{n} \int_0^1 \mathbf{1} \left\{ |W_n(r)| \leq \frac{1}{\sqrt{n}} \right\} dr &= \sqrt{n} \int_0^1 \mathbf{1} \left\{ |W(r)| \leq \frac{1}{\sqrt{n}} \right\} dr + o_p(1) \\ &= 2L(1, 0) + o_p(1), \end{aligned}$$

and thus for any  $\delta > 0$

$$n^{-1/2+\delta} \sum_{t=1}^n |x_t|^\kappa \xrightarrow{p} \infty,$$

thereby establishing the stated result.  $\square$

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