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ASYMPTOTICS FOR LINEAR PROCESSES

by

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0. ABSTRACT

A method of deriving asymptotics for linear processes is introduced which uses an explicit algebraic decomposition of the linear filter. The method leads to substantial simplifications in the asymptotics and offers a unified approach to strong laws and central limit theory for linear processes. Sample means and sample covariances are covered. The results also accommodate both homogeneous and heterogeneous innovations as well as innovations with undefined means and variances.

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1. INTRODUCTION

Since the work of McLeish (1975a, 1975b, 1977) a popular approach to the development of asymptotics for time series has been the use of limit theorems for dependent random variables that satisfy certain mixing conditions. This approach has the advantage of allowing for heterogeneity as well as dependence; it highlights the trade off that occurs in limit theory between moment conditions that control outlier probabilities and memory conditions that control the extent of the temporal dependence; and it conveniently accommodates nonlinear functional dependence on a series' past history. The latter property has ensured that the method is especially popular in the development of asymptotics for nonlinear statistical models (e.g. Gallant (1987), Ch. 7).

In spite of these advantages the approach does have some drawbacks. First, not all linear processes are strong mixing, for example, and it is necessary to use functions of mixing processes to accommodate even simple time series like the first order autoregression in a general theory. This is unfortunate because most of the stationary time series literature is still concerned with parametric models that fall in the linear process class. Second, the mixingale theory of McLeish is articulated in the L_2 norm and is therefore inapplicable in time series models with infinite variance errors.

The aim of the present paper is to show the versatility of an alternative approach that is especially designed for linear processes. In this sense, the paper represents something of a return to more traditional methods and models such as those emphasized in major textbooks like Anderson (1971), Fuller (1976) and Hannan (1970), all of which put linear processes in a central position in the development of time series asymptotics. Our method involves little in the way of probabilistic sophistication and relies almost exclusively on limit theory for independent and identically distributed (i.i.d.) or independent and non identically distributed (i.n.i.d.) random variables (r.v.'s) and for martingale difference sequences (m.d.s.'s). The key to the approach is an algebraic decomposition of the

linear filter into long-run and transitory elements that is known in the econometric literature as the Beveridge–Nelson or BN decomposition—see Beveridge and Nelson (1981). The long-run component in this decomposition yields the martingale approximation to the partial sum process of a stationary time series. In this way, the approach is related to the method pioneered by Gordin (1969) for developing central limit theorem (CLT's) for stationary processes via corresponding results for approximating martingales. A detailed treatment of that method is given in Chapter 5 of Hall and Heyde (1980), a reference we shall cite frequently as H^2 .

Since our own approach relies on a purely algebraic decomposition of the linear filter it has some advantages over the martingale approximation approach. First, it can be readily used when the time series innovations are heterogeneously distributed rather than stationary and ergodic martingale differences. Second, we may relax moment conditions and work with innovations whose first and second moments are not finite. Our method can therefore accommodate a limit theory for moving averages of r.v.'s with regularly varying tail probabilities, such as that developed in recent work by Davis and Resnick (1985a, 1985b, 1986).

A wide spectrum of limit results are presented. We give strong laws of large numbers (SLLN's), CLT's and invariance principles (IP's), we include sample means and sample covariances of stationary and nonstationary time series and we give stable limit laws for sample moments of linear processes whose domain of attraction is not the normal distribution. Few of the results given are new and our main purpose is to exhibit a unifying theme in the treatment of linear process asymptotics. The approach should be of pedagogical interest to time series specialists.

2. PRELIMINARIES

We start with a simple polynomial decomposition that is fundamental to our approach.

2.1. LEMMA (BN)

Let $C(L) = \sum_0^\infty c_j L^j$. Then

$$(1) \quad C(L) = C(1) - (1-L)\tilde{C}(L)$$

where $\tilde{C}(L) = \sum_0^\infty \tilde{c}_j L^j$, $\tilde{c}_j = \sum_{j+1}^\infty c_k$. If $p \geq 1$ then

$$(2) \quad \sum_1^\infty j^p |c_j|^p < \infty \Rightarrow \sum_0^\infty |\tilde{c}_j|^p < \infty.$$

If $p < 1$ then

$$(3) \quad \sum_1^\infty j |c_j|^p < \infty \Rightarrow \sum_0^\infty |\tilde{c}_j|^p < \infty. \quad \square$$

2.2. REMARKS

(i) For linear processes such as (13) below the decomposition (1) yields directly the martingale approximation to the partial sum process of a stationary time series (see H^2 Ch. 5). Since (1) is purely algebraic it turns out to be a useful device in reducing time series asymptotics to known theorems for i.i.d., i.n.i.d. and m.d.s. sequences. The decomposition can also be applied to deduce asymptotics for higher order moments, invariance principles and stable limit laws for time series.

(ii) When $p = 2$ we have $\sum_0^\infty \tilde{c}_j^2 < \infty$ under

$$(S_1) \quad \sum_1^\infty j^2 c_j^2 < \infty.$$

This is weaker than the commonly occurring condition

$$(S_2) \quad \sum_1^\infty j^{1/2} |c_j| < \infty.$$

Observe that $c_j = 1/j^{3/2} \ln(j+1)$, for example, satisfies (S_1) but fails (S_2) .

(iii) The algebraic decomposition (1) was used explicitly (but without conditions on the coefficients) by Beveridge and Nelson (1981) to decompose aggregate economic time series into permanent and transitory components. For convenience we shall refer to (1) subsequently as the BN decomposition although it must certainly have been known and used in earlier work. For finite lag polynomials the decomposition was used by Fuller (1976, p. 374) and by Bewley (1979). A proof of the result under (S_2) was given in Solo (1989). \square

For later development it will be useful to have available some standard asymptotics for sequences of independent r.v.'s and martingale differences. We start with the following result of Heyde and Seneta (see H^2 p. 36).

2.3. THEOREM (LLN). Let (Z_n) be a sequence of r.v.'s adapted to the filtration (\mathcal{F}_n) . Let Z be a dominating r.v. for which $E|Z| < \infty$ and

$$(4) \quad P(|Z_n| \geq x) \leq cP(|Z| \geq x)$$

for each $x \geq 0$, $n \geq 1$ and for some constant c . Then as $n \rightarrow \infty$

$$(5) \quad n^{-1} \sum_{i=1}^n [Z_i - E(Z_i | \mathcal{F}_{i-1})] \xrightarrow{p} 0.$$

If $E(|Z| \ln^+ |Z|) < \infty$ or if the Z_n are independent or if (Z_n) is stationary and \mathcal{F}_n is the natural filtration of Z_n then a.s. convergence applies in (5). \square

2.4. REMARKS

(i) A sequence (Z_n) satisfying (4) is said to be *strongly uniformly integrable* (s.u.i.)—see Billingsley (1968, p. 32) and Solo (1982/86).

(ii) If the Z_n are identically distributed with $E|Z_0| < \infty$ then (4) is automatic and (5) holds with a.s. convergence. \square

For our central limit theory a useful starting point is the following result of McLeish (1974) (see H² p. 58). Suppose (Z_i, \mathcal{F}_i) is an m.d.s., $S_n = \sum_1^n Z_i$, $U_n^2 = \sum_1^n Z_i^2$ and $s_n^2 = E(U_n^2) = E(S_n^2)$. Then

2.6. THEOREM (CLT). *If (6) and (7) hold then $s_n^{-1}S_n \rightarrow_d N(0,1)$:*

$$(6) \quad s_n^{-2}U_n^2 \rightarrow_p 1,$$

$$(7) \quad \max_{1 \leq i \leq n} |Z_{ni}| \rightarrow_p 0, \quad Z_{ni} = s_n^{-1}Z_i. \quad \square$$

The invariance principle calls for more notation. Let $[nr]$ denote the integer part of nr with $0 \leq r \leq 1$ and set

$$W_n(r) = s_n^{-1}S_{[nr]},$$

$$\xi_n(r) = s_n^{-1}S_i + s_n^{-1}Z_{i+1}(s_{i+1}^2 - s_i^2)^{-1}(rs_n^2 - s_i^2) \text{ for } s_i^2 \leq rs_n^2 < s_{i+1}^2.$$

From Brown (1971) (see H² p. 99) we have:

2.7. THEOREM (IP). *If (6) and either (7) or (8) hold then $W_n(r), \xi_n(r) \rightarrow_d W(r)$, a standard Brownian motion on $C[0,1]$, where*

$$(8) \quad \sum_1^n E[Z_{ni}^2 1(|Z_{ni}| > \epsilon)] \rightarrow 0,$$

for any $\epsilon > 0$. In fact, when (6) holds, conditions (7) and (8) are equivalent.

2.8. ASSUMPTIONS

We work with two classes of assumptions concerning the time series innovations when these have finite means. They are letter coded as: "A" for homogeneity assumptions; and "B" for heterogeneity assumptions.

$$(A_1) \quad (\epsilon_t) \text{ is i.i.d. with zero mean and } E|\epsilon_0| < \infty. \quad \square$$

(A₂) (ϵ_t) is i.i.d. with zero mean and $E(\epsilon_0^2) < \infty$.

(A₃) (ϵ_t) is i.i.d. with zero mean and finite fourth cumulant κ_4 .

(B₁) (ϵ_t) is an m.d.s. and is s.u.i. with dominating r.v. Z that satisfies $E(|Z| \ln^+ |Z|) < \infty$.

(B₂) (ϵ_t) is an m.d.s., is s.u.i. with $E(Z^2 \ln^+ |Z|) < \infty$ and further $n^{-1} \sum_1^n E(\epsilon_t^2 | \mathcal{F}_{t-1}) \xrightarrow{\text{a.s.}} \sigma_\epsilon^2$.

For the case where the innovations ϵ_t may have undefined means we make the following domain of attraction assumptions (letter coded as "C"). We say that ϵ is in the domain of attraction of a stable law with a parameter α and write $\epsilon \in \mathcal{D}(\alpha)$ if

$$(9) \quad P(\epsilon > x) = c_1 x^{-\alpha} L(x) (1 + \alpha_1(x)), \quad x > 0, \quad c_1 \geq 0$$

and

$$(10) \quad P(\epsilon < -x) = c_2 x^{-\alpha} L(x) (1 + \alpha_2(x)), \quad x > 0, \quad c_2 \geq 0$$

with $0 < \alpha < 2$, $L(x)$ a slowly varying function at ∞ and $\alpha_i(x) \rightarrow 0$ as $|x| \rightarrow \infty$. If $L(x) = 1$ in (9) and (10), then ϵ is in the normal domain of attraction of a stable law with parameter α and we write $\epsilon \in \mathcal{ND}(\alpha)$.

(C₁) (ϵ_t) is i.i.d. and $\epsilon_t \in \mathcal{D}(\alpha)$. If $\alpha > 1$, $E(\epsilon_t) = 0$ and if $\alpha = 1$ then $\epsilon_t \stackrel{d}{=} -\epsilon_t$ (i.e. ϵ_t is symmetrically distributed).

(C₂) (ϵ_t) is i.i.d. and $\epsilon_t \in \mathcal{ND}(\alpha)$. If $\alpha > 1$, $E(\epsilon_t) = 0$ and if $\alpha = 1$ then $\epsilon_t \stackrel{d}{=} -\epsilon_t$.

2.9. REMARKS

(i) It follows from Theorem LLN that under (\mathcal{A}_1) or (\mathcal{A}_2) or (\mathcal{B}_1) we have $n^{-1}\sum_1^n \epsilon_t \rightarrow_{\text{a.s.}} 0$.

(ii) Similarly under (\mathcal{B}_2) we have both $n^{-1}\sum_1^n \epsilon_t \rightarrow_{\text{a.s.}} 0$ and $n^{-1}\sum_1^n \epsilon_t^2 \rightarrow_{\text{a.s.}} \sigma_\epsilon^2$.

(iii) For (\mathcal{C}_1) and (\mathcal{C}_2) we define the normalizing sequence

$$a_n = \inf\{x : P(|\epsilon_0| > x) \leq n^{-1}\}.$$

Under (\mathcal{C}_1) we have $a_n = n^{1/\alpha} L'(n)$ where $L'(n)$ is slowly varying at infinity. Under (\mathcal{C}_2) we have $a_n = cn^{1/\alpha}$ for some constant c ; when $\epsilon_t =_d -\epsilon_t$ and $c_1 = c_2 = a^\alpha$ in (9) and (10) then $c = a$. With this construction we have the following results under either (\mathcal{C}_1) or (\mathcal{C}_2) and $0 < \alpha < 2$:

$$(11) \quad a_n^{-1}\sum_1^n \epsilon_t \rightarrow_d U_\alpha(1), \quad a_n^{-1}\sum_1^{[nr]} \epsilon_t \rightarrow_d U_\alpha(r),$$

$$(12) \quad (a_n^{-1}\sum_1^{[nr]} \epsilon_t, a_n^{-2}\sum_1^{[nr]} \epsilon_t^2) \rightarrow_d (U_\alpha(r), \int_0^r (dU_\alpha)^2).$$

Here $U_\alpha(r)$ is the Levy α -stable process and $\int_0^r (dU_\alpha)^2 = [U_\alpha]_r$ is its quadratic variation process. The first result of (11) is classical (e.g. Ibragimov and Linnik (1971), Ch. 2); the second is its functional version; and (12) is a joint functional limit law for first and second sample moments that is proved in Resnick (1986, pp. 94–95).

3. LIMIT THEORY FOR LINEAR PROCESSES

3(a) BN in direct mode and homogeneous innovations

Suppose X_t is the linear process

$$(13) \quad X_t = C(L)\epsilon_t = \sum_0^\infty c_j \epsilon_{t-j}, \quad C(L) = \sum_0^\infty c_j L^j$$

with

$$(14) \quad \sum_0^\infty c_j^2 < \infty.$$

Our object is to show how simply some of the classical time series asymptotics can be worked out by applying the BN decomposition (1) directly to (13). For (ϵ_t) we employ either (\mathcal{A}_1) or (\mathcal{A}_2) . Applying (1) to (13) we get

$$(15) \quad X_t = C(1)\epsilon_t + \tilde{\epsilon}_{t-1} - \tilde{\epsilon}_t$$

with

$$\tilde{\epsilon}_t = \tilde{C}(L)\epsilon_t = \sum_0^{\infty} \tilde{c}_j \epsilon_{t-j}, \quad \tilde{c}_j = \sum_{j+1}^{\infty} c_k.$$

Under (\mathcal{A}_2) , $E(\tilde{\epsilon}_t^2) < \infty$ if

$$(16) \quad \sum_0^{\infty} \tilde{c}_j^2 < \infty,$$

which by the BN lemma holds if (\mathcal{S}_1) holds. Now sum (15) to find

$$(17) \quad n^{-1} \sum_1^n X_t = C(1)n^{-1} \sum_1^n \epsilon_t + n^{-1}(\tilde{\epsilon}_0 - \tilde{\epsilon}_n).$$

So a SLLN for X_t follows directly from a SLLN for ϵ_t (see Remark 2.9(i)) if only

$$n^{-1}\tilde{\epsilon}_0, \quad n^{-1}\tilde{\epsilon}_n \xrightarrow{\text{a.s.}} 0.$$

These hold if

$$\sum_1^{\infty} E(\tilde{\epsilon}_0^2)n^{-2}, \quad \sum_1^{\infty} E(\tilde{\epsilon}_n^2)n^{-2} < \infty$$

which hold if $E(\tilde{\epsilon}_0^2) = E(\tilde{\epsilon}_n^2) < \infty$, which holds if (\mathcal{S}_1) does. Thus, we have established:

3.1. THEOREM (SLLN). *Under (\mathcal{A}_2) and (\mathcal{S}_1) , $n^{-1} \sum_1^n X_t \xrightarrow{\text{a.s.}} 0$. \square*

With a little more effort and a strengthening of (\mathcal{S}_1) we can relax the second moment condition in (\mathcal{A}_2) , giving:

3.2. THEOREM (SLLN). *Under (\mathcal{A}_1) and (\mathcal{S}_3) , $n^{-1} \sum_1^n X_t \xrightarrow{\text{a.s.}} 0$, where*

$$(\mathcal{S}_3) \quad \sum_1^{\infty} j|c_j| < \infty. \quad \square$$

Continuing with quick results, we now use the BN decomposition to deliver a CLT and IP for partial sums of X_t . From (15)

$$(18) \quad n^{-1/2} \sum_1^{[nr]} X_t = C(1) n^{-1/2} \sum_1^{[nr]} \epsilon_t + n^{-1/2} \tilde{\epsilon}_0 - n^{-1/2} \tilde{\epsilon}_{[nr]}.$$

Under (\mathcal{A}_2) we easily obtain a CLT and IP for the first term from Theorem 2.7. We need

$$n^{-1} \sum_1^n \epsilon_t^2 \rightarrow_p \sigma_\epsilon^2$$

which holds by Theorem LLN; this ensures that (6) holds and (8) follows because

$$n^{-1} \sum_1^n E[\epsilon_i^2 1(\epsilon_i^2 > n\delta)] = E[\epsilon_0^2 1(\epsilon_0^2 > n\delta)] \rightarrow 0, \text{ for any } \delta > 0$$

by dominated convergence. Thus, $n^{-1/2} \sum_1^n \epsilon_t \rightarrow_d N(0, \sigma_\epsilon^2)$ and $n^{-1/2} \sum_1^{[nr]} \epsilon_t \rightarrow_d \sigma_\epsilon W(r)$ by Theorems CLT and IP. To prove the CLT for X_t we see from (18) with $r = 1$ that it is sufficient to show that

$$n^{-1} \tilde{\epsilon}_0^2, n^{-1} \tilde{\epsilon}_n^2 \rightarrow_p 0.$$

These hold if $E(\tilde{\epsilon}_n^2) = E(\tilde{\epsilon}_0^2) < \infty$ which as before hold if (S_1) does.

For the IP for X_t we need (Billingsley (1968), Theorem 4.1):

$$(19) \quad \sup_r |n^{-1/2} \sum_1^{[nr]} X_t - C(1) n^{-1/2} \sum_1^{[nr]} \epsilon_t| \leq n^{-1/2} \tilde{\epsilon}_0 + \sup_r |n^{-1/2} \tilde{\epsilon}_{[nr]}| \rightarrow_p 0$$

which holds if

$$\max_{1 \leq k \leq n} (n^{-1} \tilde{\epsilon}_k^2) \rightarrow_p 0.$$

This is equivalent to

$$(20) \quad J_n = n^{-1} \sum_1^n [\tilde{\epsilon}_k^2 1(\tilde{\epsilon}_k^2 > n\delta)] \rightarrow_p 0, \text{ for any } \delta > 0$$

(c.f. H^2 p. 53 and (36) below). But (20) holds because

$$E(J_n) = E[\tilde{\epsilon}_0^2 1(\tilde{\epsilon}_0^2 > n\delta)] \rightarrow 0$$

by dominated convergence since (S_1) ensures that $E(\tilde{\epsilon}_0^2) < \infty$. We therefore have the following:

3.3. THEOREM (*CLT and IP for means*). Under (A_2) and (S_1)

$$(a) \quad n^{-1/2} \sum_1^n X_t \rightarrow_d N(0, \sigma_\epsilon^2 C(1)^2),$$

$$(b) \quad n^{-1/2} \sum_1^{[nr]} X_t \rightarrow_d \sigma_\epsilon W(r). \quad \square$$

3.4. REMARKS

(i) We have used (S_1) as the summability condition in Theorem 3.3 but it is clear from the proof that the results hold under (16), which is precisely the condition given for the use of the IP in H^2 (Theorem 5.5, p. 141 and 146) due to Heyde (1975). (S_1) may be preferable for applications because it is a little more concrete in terms of the coefficients of the process (13).

(ii) One advantage of Theorem 3.3(b) is that no proof of tightness for the partial sums of dependent sequences is required. Under (19) all we need to call upon is the IP for partial sums of i.i.d. sequences and here we can rely on existing tightness arguments with no difficulty (e.g. Billingsley (1968), pp. 137–138). \square

Our next step is to use a second order BN decomposition to establish the limit theory for sample variances. We start by writing

$$(21) \quad X_t^2 = (C(L)\epsilon_t)^2 = X_{at} + 2X_{bt}$$

with

$$\begin{aligned}
X_{at} &= \sum_0^{\infty} c_j^2 \epsilon_{t-j}^2 = f_0(L) \epsilon_t^2 \\
X_{bt} &= \sum_{r=1}^{\infty} \sum_{j=0}^{\infty} c_j c_{j+r} \epsilon_{t-j} \epsilon_{t-j-r} = \sum_1^{\infty} f_r(L) \epsilon_t \epsilon_{t-r} \\
(22) \quad f_j(L) &= \sum_{k=0}^{\infty} c_k c_{k+j} L^k = \sum_{k=0}^{\infty} f_{jk} L^k .
\end{aligned}$$

Next, employ the BN decomposition to the lag polynomial $f_j(L)$ giving

$$(23) \quad f_j(L) = f_j(1) - (1-L)\tilde{f}_j(L)$$

with

$$\tilde{f}_j(L) = \sum_{k=0}^{\infty} \tilde{f}_{jk} L^k, \quad \tilde{f}_{jk} = \sum_{s=k+1}^{\infty} f_{js} = \sum_{s=k+1}^{\infty} c_s c_{s+j} .$$

The validity of (23) follows from

3.5. LEMMA

$$(a) \quad \sum_{k=0}^{\infty} \tilde{f}_{jk}^2 = \sum_{k=0}^{\infty} (\sum_{s=k+1}^{\infty} c_s c_{s+j})^2 < \infty$$

and

$$(b) \quad \sum_{j=0}^{\infty} (\sum_{s=0}^{\infty} c_s c_{s+j})^2 < \infty$$

if

$$(S_4) \quad \sum_1^{\infty} s^{1/2} c_s^2 < \infty . \quad \square$$

We use the decomposition (22) on both components of (21), viz.

$$(24) \quad X_{at} = f_0(1) \epsilon_t^2 - (1-L)\tilde{X}_{at}$$

$$(25) \quad X_{bt} = \epsilon_t \epsilon_{t-1}^f - (1-L)\tilde{X}_{bt}$$

where

$$\begin{aligned}
\tilde{X}_{at} &= \tilde{f}_0(L)\epsilon_t^2; \quad \tilde{f}_0(L) = \sum_0^{\infty} \tilde{f}_{0k} L^k, \quad \tilde{f}_{0k} = \sum_{k+1}^{\infty} f_{0s} = \sum_{k+1}^{\infty} c_s^2 \\
(26) \quad \epsilon_{t-1}^f &= \sum_1^{\infty} f_j(1)\epsilon_{t-j} = \sum_1^{\infty} \bar{\gamma}_j \epsilon_{t-j} \\
\tilde{X}_{bt} &= \sum_1^{\infty} \tilde{f}_j(L)\epsilon_t \epsilon_{t-j}.
\end{aligned}$$

Observe that $\bar{\gamma}_j = f_j(1) = \sum_0^{\infty} c_{s+j}$ and the autocovariance function of X_t is

$$\gamma_j = E(X_0 X_j) = \sigma_{\epsilon}^2 \bar{\gamma}_j.$$

Finally, under (S_4) by Lemma 3.5(b) we have:

$$(27) \quad \sigma_f^2 = E(\epsilon_{t-1}^f)^2 = \sigma_{\epsilon}^2 \sum_1^{\infty} \bar{\gamma}_j^2 < \infty.$$

As we did for the sample mean, the approach is now to develop a SLLN and a CLT for partial sums of X_t^2 by summing in (21), (24) and (25), using results for the innovations ϵ_t and disposing of the terms that involve \tilde{X}_a and \tilde{X}_b . We obtain:

3.6. THEOREM (*SLLN for variances*). Under (A_2) and (S_5) ,

$$n^{-1} \sum_1^n X_t^2 \xrightarrow{\text{a.s.}} \gamma_0 = E(X_0^2) = \sigma_{\epsilon}^2 \sum_0^{\infty} c_s^2$$

where

$$(S_5) \quad \sum_1^{\infty} s c_s^2 < \infty. \quad \square$$

3.7. THEOREM (*CLT and IP for variances*). Under (A_3) and (S_5) ,

$$(a) \quad n^{-1/2} \sum_1^n (X_t^2 - \gamma_0) \rightarrow_d N(0, v(0))$$

where $v(0) = (2\sigma_{\epsilon}^4 + \kappa_4) f_0(1)^2 + 4\sigma_{\epsilon}^2 \sigma_f^2 = \kappa_4 \bar{\gamma}_0^2 + 2\sigma_{\epsilon}^4 \sum_{-\infty}^{\infty} \bar{\gamma}_j^2$.

$$(b) \quad n^{-1/2} \sum_1^{\lfloor nr \rfloor} (X_t^2 - \gamma_0) \rightarrow_d v(0)^{1/2} W(r). \quad \square$$

3.8. REMARKS

(i) Sample covariances may be treated in the same way as variances by using a second order BN decomposition. We write

$$\begin{aligned}
X_t X_{t+h} &= C(L)\epsilon_t C(L)\epsilon_{t+h} \\
&= \sum_0^{\infty} c_j c_{j+h} \epsilon_{t-j}^2 + \sum_{j=0}^{\infty} \sum_{r=-h-j, \neq 0}^{\infty} c_j c_{h+j+r} \epsilon_{t-j} \epsilon_{t-j-r} \\
&= f_h(L)\epsilon_t^2 + \sum_{r=-\infty, \neq 0}^{\infty} \sum_{j=-h-r}^{\infty} c_j c_{h+j+r} \epsilon_{t-j} \epsilon_{t-j-r} \\
&= f_h(L)\epsilon_t^2 + \sum_{r=-\infty, \neq 0}^{\infty} f_{h+r}(L)\epsilon_{t+h-r} \epsilon_{t+h} \\
&= f_h(L)\epsilon_t^2 + \sum_{r=1}^{\infty} [f_{h+r}(L)\epsilon_{t+h-r} \epsilon_{t+h} + f_{h-r}(L)\epsilon_{t+h+r} \epsilon_{t+h}] \\
(28) \quad &= f_h(1)\epsilon_t^2 + \sum_{r=1}^{\infty} [f_{h+r}(1)\epsilon_{t+h-r} \epsilon_{t+h} + f_{h-r}(1)\epsilon_{t+h+r} \epsilon_{t+h}] \\
&\quad - (1-L)\tilde{f}_h(L)\epsilon_t^2 - (1-L)\sum_{r=1}^{\infty} [\tilde{f}_{h+r}(L)\epsilon_{t+h-r} \epsilon_{t+h} + \tilde{f}_{h-r}(L)\epsilon_{t+h+r} \epsilon_{t+h}].
\end{aligned}$$

Without detailing all the remainder algebra we now get, as in Theorem 3.6,

$$n^{-1} \sum_1^n X_t X_{t+h} \rightarrow_{\text{a.s.}} f_h(1)\sigma_\epsilon^2 = \gamma_h;$$

and, as in Theorem 3.7, we have

$$\begin{aligned}
n^{-1/2} \sum_1^n (X_t X_{t+h} - \gamma_h) &\sim f_h(1)[n^{-1/2} \sum_1^n (\epsilon_t^2 - \sigma_\epsilon^2)] \\
&\quad + \sum_{r=1}^{\infty} [f_{h+r}(1) + f_{h-r}(1)][n^{-1/2} \sum_1^n \epsilon_t \epsilon_{t-r}] \\
&\rightarrow_d N(0, v(h)),
\end{aligned}$$

with

$$\begin{aligned}
v(h) &= f_h(1)^2 (2\sigma_\epsilon^4 + \kappa_4) + \sum_{r=1}^{\infty} (f_{h+r}(1) + f_{h-r}(1))^2 \sigma_\epsilon^4 \\
&= \kappa_4 \bar{\gamma}_h^2 + \sigma_\epsilon^4 \sum_{r=-\infty}^{\infty} [\bar{\gamma}_{h+r}^2 + \bar{\gamma}_{h+r} \bar{\gamma}_{h-r}].
\end{aligned}$$

(ii) Results for sample correlations also follow easily. Set

$$r_h = \left[\sum_1^n X_t^2 \right]^{-1} \sum_1^n X_t X_{t+h}, \quad \rho_h = \gamma_0^{-1} \gamma_h = f_0(1)^{-1} f_h(1).$$

Then, using (28), we obtain

$$(29) \quad n^{1/2}(r_h - \rho_h) \sim \left[n^{-1} \sum_1^n X_t^2 \right]^{-1} \left\{ \sum_{r=1}^{\infty} [f_{h+r}(1) + f_{h-r}(1) - \rho_h(f_r(1) + f_{-r}(1))] n^{-1/2} \sum_1^n \epsilon_t \epsilon_{t-r} \right\} \\ \rightarrow_d N(0, w(h))$$

with

$$w(h) = \sum_{r=1}^{\infty} (\rho_{h+r} + \rho_{h-r} - 2\rho_h \rho_r)^2.$$

The result for the limit distribution of serial correlations holds as in Theorem 3.7. But, in view of (29), we need only (\mathcal{A}_2) rather than (\mathcal{A}_3) , thereby corresponding to the original result of Anderson and Walker (1964)—see Anderson (1971, p. 489) and H^2 (p. 188). \square

The results in this section are not, in general, the best possible. But the approach has the advantage that the results come very easily, it involves just algebraic calculation on top of i.i.d. limit theory, and the role of the summability conditions on the coefficients of the linear process is easily understood. For these reasons the approach seems to be quite useful for pedagogical purposes.

The price we pay for the convenience of the explicit use of the BN decomposition lies in the summability conditions that are employed in its justification. To obtain improved results we weaken these conditions and use the BN decomposition only indirectly, as we now demonstrate.

3(b) BN in indirect mode and homogeneous innovations

The idea behind the indirect approach is to use the BN decomposition to suggest an appropriate approximation and then to analyze the error of approximation rather than work directly with the remainder terms in the BN construction. Thus, in the case of the linear process (13) the BN decomposition gives $C(1)\epsilon_t$ as an approximation to X_t . Instead of working with $\tilde{\epsilon}_t$ as in the explicit construction (15) we now consider the remainder $X_t - C(1)\epsilon_t$. For the CLT we have

$$n^{-1/2}w_n = n^{-1/2}\Sigma_1^n[X_t - C(1)\epsilon_t]$$

and the CLT follows if we can show $n^{-1}E(w_n^2) \rightarrow 0$. Similarly, the SLLN follows if we can show $n^{-1}w_n \xrightarrow{\text{a.s.}} 0$. This approach has been used before and is evident, for example, in Hannan (1970, pp. 246–248) and in H^2 (Theorem 5.3, pp. 133–134) in the proof of time series CLT's.

The following results make systematic use of the method.

3.9. THEOREM (SLLN). Under (\mathcal{A}_2) and (\mathcal{S}_6) , $n^{-1}\Sigma_1^n X_t \xrightarrow{\text{a.s.}} 0$, where

$$(\mathcal{S}_6) \quad \Sigma_1^\infty \ln s |c_s| < \infty. \quad \square$$

3.10. THEOREM (CLT). Under (\mathcal{A}_2) and (\mathcal{S}_7) , $n^{-1/2}\Sigma_1^n X_t \xrightarrow{d} N(0, \sigma_\epsilon^2 C(1)^2)$, where

$$(\mathcal{S}_7) \quad 0 < |C(1)| < \infty \quad \square$$

3.11. REMARKS

(i) A result that is very similar to Theorem 3.9 is obtained by using McLeish's (1975a) mixingale convergence theorem (H^2 , p. 22 and p. 41). Let $\mathcal{F}_t = \sigma(\epsilon_t, \epsilon_{t-1}, \dots)$ be the natural filtration for (ϵ_t) and set

$$\psi_m = \|E(X_t | \mathcal{F}_{t-m})\|_2 = \left\{ E[E(X_t | \mathcal{F}_{t-m})^2] \right\}^{1/2} = \left[\Sigma_m^\infty c_s^2 \right]^{1/2} \sigma_\epsilon.$$

Then, McLeish's SLLN (Corollary 1.9 and Example 1 of McLeish (1975a)) requires ψ_m to be of size $-1/2$, i.e. $\psi_m = O(m^{-1/2}L(m)^{-1})$, where $L(m)$ is slowly varying at ∞ and satisfies the summability requirement $\sum_1^{\infty} m^{-1}L(m)^{-1} < \infty$. This leads to $\psi_m^2 = O(m^{-1}L(m)^{-2})$ and since $\sum_1^{\infty} m^{-1}L(m)^{-2} < \infty$ we deduce the implied summability condition

$$\sum_{m=1}^{\infty} \sum_{s=m}^{\infty} c_s^2 = \sum_0^{\infty} s c_s^2 < \infty.$$

This is our (S_5) and is weaker than (S_6) , but only by a slowly varying factor. For example, $c_2 = s^{-1}[\ln(1+s)]^{-1}$ satisfies (S_5) but fails (S_6) . Thus McLeish's mixingale approach leads to a stronger result but involves more work and sophistication.

(ii) Theorem 3.10 offers a new proof of the minimal result given in H^2 (Corollary 5.2, p. 135)—see also Hannan (1970, Theorem 11, p. 221). \square

3(c) BN in direct mode and heterogeneous innovations

The explicit form of (15) makes it just as easy to work with heterogeneous as homogeneous innovations in the decomposition. Again the simplicity of the direct mode approach is that we can appeal immediately to established theory for i.n.i.d. and m.d.s. sequences and need only attend to the remainder terms to produce a rigorous theory for linear processes under assumptions like (B_1) or (B_2) on the innovations. The following are a selection of first and second moment results that are easy to obtain. The proofs are just like those for the homogeneous case.

3.12. THEOREM (SLLN). Under (B_2) and (S_1) , $n^{-1}\sum_1^n X_t \rightarrow_{\text{a.s.}} 0$. \square

3.13. THEOREM (SLLN). Under (B_1) and (S_3) , $n^{-1}\sum_1^n X_t \rightarrow_{\text{a.s.}} 0$. \square

3.14. THEOREM (CLT and IP). Under (B_2) and (S_1) ,

$$(a) \quad n^{-1/2} \sum_1^n X_t \rightarrow_d N(0, \sigma_\epsilon^2 C(1)^2).$$

Under (S_3) and with (B_2) strengthened so that $E(Z^{2+\eta}) < \infty$ for some $\eta > 0$,

$$(b) \quad n^{-1/2} \sum_1^{[nr]} X_t \rightarrow_d \sigma_\epsilon C(1) W(r). \quad \square$$

3.15. THEOREM (SLLN for variances). Under (S_5) and with (B_2) strengthened so that $E(Z^4) < \infty$, we have $n^{-1} \sum_1^n X_t^2 \rightarrow_{a.s.} \gamma_0$. \square

3.16. REMARKS

(i) Theorem 3.12 is related to a result of Hannan and Heyde (1972) (H^2 p. 184) who require only

$$(S_8) \quad \sum_0^\infty |c_s| < \infty$$

in place of (S_1) and only $E(Z^2)$ in (B_2) .

(ii) Theorem 3.13 gives us an extension of the Markov SLLN to linear processes. The theorem continues to hold, by virtue of theorem LLN, if we replace (B_1) with:

$$(B_1)' \quad (\epsilon_t) \text{ is an independent sequence, is s.u.i. with dominating r.v. } Z \\ \text{and } E|Z| < \infty.$$

Thus, all that is needed to extend theorem LLN from independent sequences to linear processes is (S_3) . This result and Theorem 3.13 would appear to be new.

(iii) Theorem 3.15 is also related to Hannan and Heyde (1972) (H^2 p. 184). Again, they require only $E(Z^2) < \infty$ in our (B_2) and only (S_8) in place of our (S_1) ; but they show convergence in probability not a.s. convergence in this case.

(iv) Hannan and Heyde (1972) also extend the Anderson and Walker (1964) limit theory for autocorrelations to the heterogeneous case. Their theorem 2 may be obtained using our approach.

3.17. THEOREM (CLT for autocorrelations). If (S_8) holds and if (B_2) is strengthened so that (i) $E(Z^4) < \infty$, (ii) $E(\epsilon_t^2 \epsilon_{t-s} \epsilon_{t-r}) = \sigma^4 \tau_{rs}$ is finite and uniformly bounded for every t , $r \geq 1$, $s \geq 1$, and (iii)

$$n^{-1} \sum_{t=1}^n \epsilon_{t-r} \epsilon_{t-s} E(\epsilon_t^2 | \mathcal{F}_{t-1}) \rightarrow_{\text{a.s.}} \sigma^4 \tau_{rs}$$

as $n \rightarrow \infty$ uniformly for $r \geq 1$, $s \geq 1$, then we have

$$n^{1/2} (r_h - \rho_h) \rightarrow_d N(0, \mathbf{x}(h))$$

where

$$\mathbf{x}(h) = \sum_{r,s=1}^{\infty} \tau_{rs} (\rho_{h+r} + \rho_{h-r} - 2\rho_h \rho_r) (\rho_{h+s} + \rho_{h-s} - 2\rho_h \rho_s). \quad \square$$

$\mathcal{S}(d)$ BN in direct mode and stable limit laws

Stable limit laws for the linear process X_t given by (13) can be deduced in much the same way as the classical SLLN and CLT asymptotics. We rely again on the BN decomposition that leads to (15). First observe that, if ϵ_t satisfies condition (C_1) , $X_t = \sum_0^{\infty} c_j \epsilon_{t-j}$ is convergent a.s. provided

$$(\mathcal{S}_9) \quad \sum_0^{\infty} |c_j|^p < \infty$$

for $0 < p < \alpha$ and $p \leq 1$ (e.g. Brockwell and Davis (1987), p. 480). Similarly, $\tilde{\epsilon}_t = \sum_0^{\infty} \tilde{c}_j \epsilon_{t-j}$ in (15) is convergent a.s. provided $\sum_0^{\infty} |\tilde{c}_j|^p < \infty$ and this holds by the BN lemma if

$$(\mathcal{S}_{10}) \quad \sum_1^{\infty} j |c_j|^p < \infty, \text{ for } 0 < p < \alpha \text{ and } p \leq 1.$$

With the validity of (15) in hand, it is a simple matter to deduce asymptotics for standardized sums and cross products of X_t . We give the following two useful results.

3.18. THEOREM (*stable limit for means*). Under (C_1) and (S_{10}) , $a_n^{-1} \sum_1^n X_t \rightarrow_d C(1)U_\alpha(1)$. \square

3.19. THEOREM (*stable limit for covariances*). Under (C_1) and (S_{10}) ,

$$(a) a_n^{-2} \sum_1^n [X_t^2, X_t X_{t+1}, \dots, X_t X_{t+h}] \rightarrow_d [f_0(1), f_1(1), \dots, f_h(1)] / \int_0^1 (dU_\alpha)^2,$$

$$(b) r_h \rightarrow_p \rho_h = f_h(1)/f_0(1). \quad \square$$

3.20. REMARKS

(i) Theorem 3.18 gives a result that seems first to have been established by Davis and Resnick (1985a, Theorem 4.1, p. 189). Their proof uses truncation arguments and point process theory and is more involved than ours; but they need only (S_9) in place of our (S_{10}) . As in section 3(a), the explicit construction (15) leads to a substantial simplification but is achieved at a greater cost.

(ii) Theorem 3.19 also gives results that appear in Davis and Resnick (1985a, Theorem 4.2, p. 192). Again, they require (S_9) rather than (S_{10}) so our results are therefore mainly of pedagogical interest.

(iii) Interestingly, Theorem 3.18 does not extend directly to a functional version, as it does in the case of finite variance innovations (cf. Theorem 3.3). This has been discovered by Avram and Taqqu (1986, 1989). In the present context, we can explain the failure in terms of the BN decomposition. We have, as before,

$$a_n^{-1} \sum_1^{[nr]} X_t = C(1) a_n^{-1} \sum_1^{[nr]} \epsilon_t + a_n^{-1} (\tilde{\epsilon}_0 - \tilde{\epsilon}_{[nr]})$$

and, according to (11), $a_n^{-1} \sum_1^{[nr]} \rightarrow_d U_\alpha(r)$ in the space $D[0,1]$ with the Skorohod topology. However, the remainder term does not vanish in probability in general. For instance, the distance between $a_n^{-1} \tilde{\epsilon}_{[nr]}$ and the zero function in the Skorohod J_1 metric is simply

$$(30) \quad \sup_r |a_n^{-1} \tilde{\epsilon}_{[nr]}| = \max_{0 \leq k \leq n} a_n^{-1} |\tilde{\epsilon}_k| .$$

But, under (S_0) , $\tilde{\epsilon}_k \in \mathcal{D}(\alpha)$ and thus, when $\alpha < 2$, (30) does not converge in probability to zero (Breiman (1965), Theorem 2, p. 323). So the functional law $a_n^{-1} \Sigma_1^{[nr]} X_t \rightarrow_d C(1)U_\alpha(r)$ does not apply in $D[0,1]$ endowed with the usual Skorohod topology, even though all the finite dimensional distributions converge. \square

4. SUPPLEMENTARY REMARKS

(i) The BN Lemma 2.1 continues to hold for matrix polynomials using conventional matrix norms in the summability conditions. Thus, the decomposition (15) also applies to vector linear processes. The limit theory of Sections 3(a)–(c) can then be extended to the multivariate case.

(ii) The decomposition (15) is important in the vector case to the theory of cointegration—see Engle and Granger (1987). Suppose $(1-L)Y_t = X_t = C(L)\epsilon_t = \sum_{j=0}^{\infty} C_j L^j \epsilon_t$ where ϵ_t is an m.d.s. and $\sum_{j=0}^{\infty} \|C_j\|^2 < \infty$ with $\|C_j\| = [\text{tr}(C_j' C_j)]^{1/2}$. Then Y_t is stationary and if $C(1) \neq 0$ at least some components of Y_t are integrated processes. The components of Y_t are cointegrated iff $C(1)$ is a singular matrix. Using (15) we have

$$(1-L)Y_t = C(1)\epsilon_t + \tilde{\epsilon}_{t-1} - \tilde{\epsilon}_t$$

and by summation and with $Y_0 = 0$ we get

$$Y_t = C(1)S_t + \tilde{\epsilon}_0 - \tilde{\epsilon}_t, \quad S_t = \sum_1^t \epsilon_s .$$

Thus, if α is a vector for which $\alpha' C(1) = 0$, then α' annihilates the integrated element $C(1)S_t$ of Y_t (i.e. the martingale approximation to Y_t) and we have

$$(31) \quad \alpha' Y_t = \alpha' (\tilde{\epsilon}_0 - \tilde{\epsilon}_t) \text{ a.s.},$$

which is a stationary time series under the stated summability condition. In econometric models, the equation (31) is interpreted as describing stationary deviations about a long-run equilibrium relation $\alpha' Y_t = 0$. Phillips (1988), Johansen (1988) and Phillips and Loretan (1989) provide further discussion and develop optimal inference procedures.

(iii) The BN lemma also has a version that is suitable for frequency domain applications. In effect, we can expand the polynomial $C(L)$ in a Taylor series about an arbitrary point on the unit circle, say $e^{i\lambda}$ rather than unity, giving

$$(32) \quad C(L) = C(e^{i\lambda}) - (1 - e^{-i\lambda}L)\tilde{C}_\lambda(L)$$

with

$$\tilde{C}_\lambda(L) = \sum_0^\infty \tilde{c}_{s\lambda} L^s, \quad \tilde{c}_{s\lambda} = e^{-i\lambda s} \sum_{s+1}^\infty c_k e^{i\lambda k}.$$

This leads to the following decomposition for the discrete Fourier transform (d.f.t.), $(2\pi n)^{-1/2} \sum_1^n X_t e^{it\lambda}$, of X_t in terms of the d.f.t. of ϵ_t and a residual:

$$w_X(\lambda) = C(e^{i\lambda})w_\epsilon(\lambda) + (2\pi n)^{-1/2} [\tilde{X}_{0\lambda} - e^{i\lambda n} \tilde{X}_{n\lambda}]$$

with

$$\tilde{X}_{t\lambda} = \tilde{C}_\lambda(L)\epsilon_t = \sum_0^\infty \tilde{c}_{s\lambda} \epsilon_{t-s}.$$

This decomposition leads in much the same way as Theorem 3.3 to a CLT for $w_X(\lambda)$ and joint CLT's for the d.f.t.'s at many frequencies. In the above form, the decomposition (32) is used by Hannan and Deistler (1988, p. 156) and is apparently due to Bouaziz. It is justified, as in Lemma BN, by conditions such as (S_1) . This frequency domain BN approach may also be applied in what we have termed the indirect mode in Section 3(b). As such, the idea appears in Hannan (1970, theorem 1, p. 248).

5. PROOFS

We start with some useful bounds provided by the following lemma.

5.1. LEMMA

- (a) $\Sigma_{t+1}^{\infty} u^{-1-b} \leq b^{-1} t^{-b}$, $b > 0$
- (b) $\Sigma_1^t u^{c-1} \leq c^{-1} t^c$, $0 < c < 1$
- (c) $\Sigma_1^t u^{-1} \leq 1 + \ell n t$.

PROOF. If $0 < u \leq s$ then $s^{-1} \leq u^{-1}$ and

$$s^{-1} \leq \int_{s-1}^s u^{-1} du = \ell n s - \ell n(s-1)$$

so that $\Sigma_{s=2}^t s^{-1} \leq \ell n t$ and (c) follows. Results (b) and (a) follow by similar arguments. \square

5.2. PROOF OF LEMMA 2.1. The case $p = 1$ is obvious so we take $p > 1$. Then for a suitably chosen constant a and by Hölder's inequality we have

$$\begin{aligned} & \Sigma_0^{\infty} \left| \Sigma_{j+1}^{\infty} c_k \right|^p \leq \Sigma_0^{\infty} \left[\Sigma_{j+1}^{\infty} k^a |c_k| k^{-a} \right]^p \\ & \leq \Sigma_0^{\infty} \left[\Sigma_{j+1}^{\infty} k^{ap} |c_k|^p \right] \left[\Sigma_{j+1}^{\infty} k^{-aq} \right]^{p/q}, \quad 1/p + 1/q = 1 \\ & \leq (aq-1)^{-1} \Sigma_0^{\infty} \left[\Sigma_{j+1}^{\infty} k^{ap} |c_k|^p \right] \left[j^{1-aq} \right]^{p/q}, \quad \text{using 5.1(a) with } 1/q < a < 1 \\ & = (aq-1)^{-1} \Sigma_1^{\infty} k^{ap} |c_k|^p \left[\Sigma_0^{k-1} j^{p/q-ap} \right] \\ & \leq \{(aq-1)(1+p/q-qp)\}^{-1} \Sigma_1^{\infty} k^{ap} |c_k|^p k^{1+p/q-ap}, \quad \text{using 5.1(b) with } a < 1/p+1/q \\ & \leq \{(aq-1)(1+p/q-ap)\}^{-1} \Sigma_1^{\infty} k^p |c_k|^p \end{aligned}$$

for $1/q < a < 1/p + 1/q = 1$. To prove (3) we note that for $p < 1$

$$\Sigma_0^{\infty} |\tilde{c}_j|^p = \Sigma_0^{\infty} \left| \Sigma_{j+1}^{\infty} c_k \right|^p \leq \Sigma_0^{\infty} \Sigma_{j+1}^{\infty} |c_k|^p = \Sigma_1^{\infty} k |c_k|^p. \quad \square$$

5.3. PROOF OF THEOREM 2.6. This follows from McLeish's (1974) theorem quoted in H^2 (p. 58). That theorem also requires $E(\max_{1 \leq i \leq n} Z_{ni}^2)$ is bounded uniformly in n . But in our context this can be dispensed with since it is bounded by $E(\sum_1^n Z_{ni}^2) = 1$. \square

5.4. PROOF OF THEOREM 2.7. This follows from Brown (1971) (see H^2 , p. 99). To show that (7) implies (8) we note that $0 \leq \sum_1^n Z_{ni}^2 1(|Z_{ni}| > \epsilon) \leq \sum_1^n Z_{ni}^2$. If each of the following are true:

$$(33) \quad \sum_1^n Z_{ni}^2 1(|Z_{ni}| > \epsilon) \xrightarrow{p} 0 \text{ for any } \epsilon > 0$$

$$(34) \quad \sum_1^n Z_{ni}^2 \xrightarrow{p} 1$$

$$(35) \quad E(\sum_1^n Z_{ni}^2) \rightarrow 1$$

then, by a version of the dominated convergence theorem (cf. H^2 , p. 281), (8) immediately follows. But (34) is simply (6), while (35) is trivially true. Finally (33) is equivalent to (8) since (cf. H^2 , p. 53)

$$(36) \quad P(\max_{1 \leq i \leq n} |Z_{ni}| > \epsilon) = P(\sum_1^n Z_{ni}^2 1(|Z_{ni}| > \epsilon) > \epsilon^2).$$

To show (8) implies (7) we have, in view of (36), only to show (33). But, of course, (8) implies (33). Thus, when (6) holds, conditions (7) and (8) are equivalent. \square

5.5. PROOF OF THEOREM 3.2. From (16)

$$(37) \quad n^{-1} \sum_1^n X_t = C(1) n^{-1} \sum_1^n \epsilon_t + n^{-1} \tilde{\epsilon}_0 - n^{-1} \tilde{\epsilon}_n.$$

By theorem LLN the first term $\xrightarrow{\text{a.s.}} 0$. Since $|\tilde{\epsilon}_0| < \infty$ a.s. the second term $\xrightarrow{\text{a.s.}} 0$ also. The third term of (37) is

$$n^{-1} \tilde{\epsilon}_n = n^{-1} \sum_0^n \tilde{c}_j \epsilon_{n-j} = n^{-1} \sum_{-\infty}^n \tilde{c}_{n-s} \epsilon_s = n^{-1} \sum_0^n \tilde{c}_{n-s} \epsilon_s + n^{-1} \sum_1^\infty \tilde{c}_{n+t} \epsilon_{-t}.$$

Introduce $\bar{c}_t = \sum_{t+1}^{\infty} |c_s|$ and note that

$$\sum_1^{\infty} |\bar{c}_t| \leq \sum_1^{\infty} \bar{c}_t = \sum_1^{\infty} \sum_{t+1}^{\infty} |c_s| \leq \sum_1^{\infty} s |c_s| < \infty$$

under (\mathcal{S}_3) . Then

$$(38) \quad \begin{aligned} n^{-1} |\tilde{\epsilon}_n| &\leq n^{-1} \sum_0^n |\bar{c}_t| |\epsilon_{n-t}| + n^{-1} \sum_1^{\infty} \bar{c}_{n+t} |\epsilon_{-t}| \\ &\leq n^{-1} \left(\max_{0 \leq t \leq n} |\epsilon_t| \right) \left(\sum_0^{\infty} |\bar{c}_t| \right) + n^{-1} \sum_1^{\infty} \bar{c}_t |\epsilon_{-t}|. \end{aligned}$$

Now $\max_{0 \leq t \leq n} (n^{-1} |\epsilon_t|) \rightarrow_{\text{a.s.}} 0$ if $n^{-1} |\epsilon_n| \rightarrow_{\text{a.s.}} 0$, which holds if $n^{-1} \epsilon_n \rightarrow_{\text{a.s.}} 0$, which holds since

$$n^{-1} \epsilon_n = n^{-1} \sum_1^n \epsilon_j - (1 - n^{-1})(n-1)^{-1} \sum_1^{n-1} \epsilon_j \rightarrow_{\text{a.s.}} 0 - 0 = 0.$$

Next $\sum_1^{\infty} \bar{c}_t |\epsilon_t| < \infty$ a.s. since its expectation is finite. Thus, both terms of (38) $\rightarrow_{\text{a.s.}} 0$ and the result follows from (37). \square

5.6. PROOF OF LEMMA 3.5.

$$\begin{aligned} \sum_{k=0}^{\infty} \left[\sum_{s=k+1}^{\infty} c_s c_{s+j} \right]^2 &= \sum_0^{\infty} \left[\sum_{k+1}^{\infty} s^{1/4} c_s c_{s+j} s^{-1/4} \right]^2 \leq \sum_0^{\infty} \left(\sum_{k+1}^{\infty} s^{1/2} c_s^2 \right) \left(\sum_{k+1}^{\infty} c_{s+j}^2 s^{-1/2} \right) \\ &\leq \left(\sum_1^{\infty} s^{1/2} c_s^2 \right) \left(\sum_{k=0}^{\infty} \sum_{s=k+1}^{\infty} s^{-1/2} c_{s+j}^2 \right) \leq \left(\sum_1^{\infty} s^{1/2} c_s^2 \right) \left(\sum_{s=1}^{\infty} s^{-1/2} c_{s+j}^2 \sum_{k=0}^{s-1} 1 \right) \\ &\leq \left(\sum_1^{\infty} s^{1/2} c_s^2 \right) \left(\sum_1^{\infty} (s+j)^{1/2} c_{s+j}^2 \right) \leq \left[\sum_1^{\infty} s^{1/2} c_s^2 \right]^2 < \infty. \end{aligned}$$

The proof of (b) follows in the same way. \square

5.7. PROOF OF THEOREM 3.6. From (21) we have to show

$$(39) \quad n^{-1} \sum_1^n X_{at} \rightarrow_{\text{a.s.}} \gamma_0$$

$$(40) \quad n^{-1} \sum_1^n X_{bt} \rightarrow_{\text{a.s.}} 0.$$

We shall prove (40) first. From (25) this follows if

$$(41) \quad n^{-1} \tilde{X}_{bn} \xrightarrow{\text{a.s.}} 0$$

and

$$(42) \quad n^{-1} \sum_1^n \epsilon_t^f \epsilon_{t-1}^f \xrightarrow{\text{a.s.}} 0.$$

By Kronecker's lemma (42) holds if $T_n = \sum_1^n t^{-1} \epsilon_t^f \epsilon_{t-1}^f$ converges a.s.. But T_n is a martingale so by the L_2 martingale convergence theorem (e.g. H², p. 18) T_n converges a.s. if $\sup_n E(T_n^2) < \infty$, which holds because $\sup_n E(T_n^2) \leq \sum_1^\infty (\sigma_\epsilon^2 \sigma_f^2 t^{-2}) < \infty$. Next (41) holds if $\sum_1^\infty E(\tilde{X}_{bn}^2) n^{-2} < \infty$, which holds if $E(\tilde{X}_{bn}^2) < \infty$, which holds if (S_5) holds, as shown in Lemma 5.8 below.

To prove (39) we note from (24) that this holds if the following are true:

$$(43) \quad n^{-1} \tilde{X}_{an} \xrightarrow{\text{a.s.}} 0$$

$$(44) \quad n^{-1} \sum_1^n \epsilon_t^2 \xrightarrow{\text{a.s.}} \sigma_\epsilon^2.$$

But (44) follows from theorem LLN under (A_2) so we need only prove (43). As in (38) above we have

$$(45) \quad n^{-1} \tilde{X}_{an} = n^{-1} \sum_0^\infty \tilde{f}_{0k} \epsilon_{n-k}^2 \leq n^{-1} \left(\max_{0 \leq t \leq n} \epsilon_t^2 \right) \left(\sum_0^\infty \tilde{f}_{0k} \right) + n^{-1} \sum_1^\infty \tilde{f}_{1k} \epsilon_{-k}^2.$$

Now $n^{-1} \max_{0 \leq t \leq n} \epsilon_t^2 \xrightarrow{\text{a.s.}} 0$ if $n^{-1} \epsilon_n^2 \xrightarrow{\text{a.s.}} 0$ which holds because

$$n^{-1} \epsilon_n^2 = n^{-1} \sum_1^n \epsilon_j^2 - (1 - n^{-1})(n-1)^{-1} \sum_1^{n-1} \epsilon_j^2 \xrightarrow{\text{a.s.}} 0.$$

Also $\sum_0^\infty \tilde{f}_{0k} = \sum_0^\infty \sum_{k+1}^\infty c_s^2 \leq \sum_0^\infty s c_s^2 < \infty$ so that the first term of (45) converges a.s. to zero. Moreover, this ensures that $E(\sum_1^\infty \tilde{f}_{0k} \epsilon_{-k}^2) < \infty$ so that $\sum_1^\infty \tilde{f}_{0k} \epsilon_{-k}^2 < \infty$ a.s. and the second term of (45) converges a.s. to zero. \square

5.8. LEMMA. Under (A_3) and (S_5) , $E(\tilde{X}_{bn}^2) < \infty$.

PROOF.

$$\begin{aligned}
E(\tilde{X}_{bn}^2) &= \sum_{j,j'=1}^{\infty} \sum_{k,k'=0}^{\infty} \tilde{f}_{jk} \tilde{f}_{j'k'} E(\epsilon_{n-k} \epsilon_{n-j-k} \epsilon_{n-k'} \epsilon_{n-j'-k'}) \\
&= \sum_{j,j'} \sum_{k,k'} \tilde{f}_{jk} \tilde{f}_{j'k'} (\delta_{jj'} \delta_{kk'} \sigma_{\epsilon}^4 + \delta_{k,k'+j'} \delta_{k',k+j} \sigma_{\epsilon}^4) \\
(46) \quad &= \sum_{j=1}^{\infty} \sum_{k=0}^{\infty} \tilde{f}_{jk}^2 \sigma_{\epsilon}^4 + \sum_{k=0}^{\infty} \sum_{k'=0}^{\infty} \tilde{f}_{k'-k,k} \tilde{f}_{k-k',k'} \sigma_{\epsilon}^4 - \sum_{k=0}^{\infty} \tilde{f}_{0k}^2 \sigma_{\epsilon}^4.
\end{aligned}$$

But the first term of (46) is bounded since

$$\begin{aligned}
(47) \quad \sum_{j=1}^{\infty} \sum_{k=0}^{\infty} \tilde{f}_{jk}^2 &= \sum_{j=1}^{\infty} \sum_{k=0}^{\infty} \left[\sum_{s=k+1}^{\infty} c_s c_{s+j} \right]^2 \\
&\leq \sum_{j=1}^{\infty} \sum_{k=0}^{\infty} (\sum_{s=k+1}^{\infty} c_s^2) (\sum_{s=k+1}^{\infty} c_{s+j}^2) \\
&\leq (\sum_{k=0}^{\infty} \sum_{s=k+1}^{\infty} c_s^2) (\sum_{j=1}^{\infty} \sum_{s=1}^{\infty} c_{s+j}^2) \\
&= \left[\sum_{s=1}^{\infty} c_s^2 \sum_{k=1}^{s-1} 1 \right]^2 \\
&= \left[\sum_1^{\infty} s c_s^2 \right]^2 < \infty.
\end{aligned}$$

Next, using the Cauchy inequality for double sums, we get for the second term

$$\begin{aligned}
(\sum_{k,k'} \tilde{f}_{k'-k,k} \tilde{f}_{k-k',k'})^2 &\leq \left[\sum_{k,k'} \tilde{f}_{k'-k,k}^2 \right]^2 \\
&= \left[\sum_{k'=0}^{\infty} \sum_{k=0}^{k'} \tilde{f}_{k'-k,k}^2 \right]^2 \\
&\leq \left[\sum_{j=1}^{\infty} \sum_{k=0}^{\infty} \tilde{f}_{jk}^2 + \tilde{f}_{00}^2 \right]^2 < \infty
\end{aligned}$$

since (47) is finite. Finally

$$\begin{aligned}
\sum_{k=0}^{\infty} \tilde{f}_{0k}^2 &= \sum_{k=0}^{\infty} \left[\sum_{s=k+1}^{\infty} c_s^2 \right]^2 \\
&\leq (\sum_1^{\infty} c_s^2) (\sum_{k=0}^{\infty} \sum_{s=k+1}^{\infty} c_s^2) \\
&\leq (\sum_1^{\infty} c_s^2) (\sum_{s=1}^{\infty} c_s^2 \sum_{k=0}^{s-1} 1) \\
&= (\sum_1^{\infty} c_s^2) (\sum_1^{\infty} s c_s^2) < \infty.
\end{aligned}$$

Thus all terms of (46) are finite and the required result follows. \square

5.9. PROOF OF THEOREM 3.7. From (21), (24) and (25) we have the decomposition

$$\begin{aligned}
(48) \quad n^{-1/2} \sum_1^n (X_t^2 - \gamma_0) &= n^{-1/2} \sum_1^n [\bar{\gamma}_0 (\epsilon_t^2 - \sigma_\epsilon^2) + 2\epsilon_t \epsilon_{t-1}^f] \\
&\quad + n^{-1/2} (\tilde{X}_{a0} - \tilde{X}_{an}) + 2n^{-1/2} (\tilde{X}_{b0} - \tilde{X}_{bn}).
\end{aligned}$$

From Lemma 5.8 $E(\tilde{X}_{bn}^2) < \infty$ and the final term of (48) converges in probability to zero.

Further

$$E(\tilde{X}_{a0}), E(\tilde{X}_{an}) = \sigma_\epsilon^2 \sum_0^{\infty} \tilde{f}_{0k} = \sigma_\epsilon^2 \sum_{k=0}^{\infty} \sum_{s=k+1}^{\infty} c_s^2 = \sigma_\epsilon^2 \sum_{s=1}^{\infty} s c_s^2 < \infty$$

so that the second term on the right side of (48) also converges in probability to zero. It remains to show that the first term of (48) converges weakly to $N(0, v(0))$. We apply theorem CLT and need only verify:

$$(49) \quad n^{-1} \sum_1^n Z_t^2 \xrightarrow{p} v(0),$$

$$(50) \quad n^{-1} \sum_1^n E[Z_t^2 1(Z_t^2 > \epsilon n)] \rightarrow 0$$

with

$$Z_t = \bar{\gamma}_0 (\epsilon_t^2 - \sigma_\epsilon^2) + 2\epsilon_t \epsilon_{t-1}^f.$$

Now (50) is just $E[Z_0^2 1(Z_0^2 > \epsilon n)] \rightarrow 0$, which follows by dominated convergence since $E(Z_0^2) = v(0) < \infty$ by (A_3) . Next we apply theorem LLN to (49) giving

$$n^{-1} \sum_1^n [Z_t^2 - E(Z_t^2 | \mathcal{F}_{t-1})] \rightarrow_p 0.$$

So we need only show that

$$n^{-1} \sum_1^n E(Z_t^2 | \mathcal{F}_{t-1}) \rightarrow_p v(0)$$

which follows directly if

$$(51) \quad n^{-1} \sum_1^n \left[\epsilon_{t-1}^f \right]^2 \rightarrow_p \sigma_f^2.$$

This holds by the pointwise ergodic theorem since $\sigma_f^2 < \infty$ but we can also establish it by appealing to our own SLLN, Theorem 3.6—see Lemma 5.10 below.

To prove the IP (part (b)) we again employ the decomposition

$$n^{-1/2} \sum_1^{[nr]} (X_t^2 - \gamma_0) = n^{-1/2} \sum_1^{[nr]} Z_t + n^{-1/2} (\tilde{X}_{a0} - \tilde{X}_{a[nr]}) + 2n^{-1/2} (\tilde{X}_{b0} - \tilde{X}_{b[nr]}).$$

Since (50) and (51) hold, we have $n^{-1/2} \sum_1^{[nr]} Z_t \rightarrow_d v(0)^{1/2} W(r)$ directly from theorem IP. The IP for $X_t^2 - \gamma_0$ then follows provided

$$\max_{1 \leq k \leq n} (n^{-1} \tilde{X}_{ak}^2) \rightarrow_p 0, \quad \max_{1 \leq k \leq n} (n^{-1} \tilde{X}_{bk}^2) \rightarrow_p 0.$$

But these are equivalent (cf. (20)) to

$$J_{an} = n^{-1} \sum_1^n [\tilde{X}_{ak}^2 1(\tilde{X}_{ak}^2 > n\delta)] \rightarrow_p 0, \quad J_{bn} = n^{-1} \sum_1^n [\tilde{X}_{bk}^2 1(\tilde{X}_{bk}^2 > n\delta)] \rightarrow_p 0$$

which in turn hold because

$$E(J_{an}) = E[\tilde{X}_{a0}^2 1(\tilde{X}_{a0}^2 > n\delta)] \rightarrow 0, \quad E(J_{bn}) = E[\tilde{X}_{b0}^2 1(\tilde{X}_{b0}^2 > n\delta)] \rightarrow 0$$

for any $\delta > 0$, by dominated convergence. \square

5.10. LEMMA. Under (\mathcal{A}_2) and (\mathcal{S}_5) , $n^{-1} \sum_1^n (\epsilon_{t-1}^f)^2 \rightarrow_{\text{a.s.}} \sigma_f^2$.

PROOF. We appeal to Theorem 3.6. Since $\epsilon_{t-1}^f = \sum_1^\infty \bar{\gamma}_s \epsilon_{t-s}$ we require $\sum_1^\infty \bar{\gamma}_s^2 < \infty$ for this theorem to be used. Observe that

$$\begin{aligned} \sum_1^\infty \bar{\gamma}_s^2 &= \sum_{s=1}^\infty \left[\sum_{t=1}^\infty c_t c_{t+s} \right]^2 \\ &< 2 \sum_{s=1}^\infty \left[\sum_{t=1}^s c_t c_{t+1} \right]^2 + 2 \sum_{s=1}^\infty \left[\sum_{t=s+1}^\infty c_t c_{t+s} \right]^2. \end{aligned}$$

The second term in the above expression is dominated by twice times

$$\begin{aligned} &\sum_{s=1}^\infty \sum_{t=s+1}^\infty c_t^2 t \sum_{t=s+1}^\infty c_{t+s}^2 t^{-1} \\ &\leq (\sum_1^\infty t c_t^2) \sum_{t=1}^\infty t^{-1} \sum_{s=1}^{t-1} s c_{s+t}^2 \\ &\leq (\sum_1^\infty t c_t^2) \sum_{t=1}^\infty \sum_{s=1}^{t-1} c_{s+t}^2 \\ &\leq (\sum_1^\infty t c_t^2) \sum_{t=1}^\infty \sum_{r=t+1}^\infty c_r^2 \\ &= \left[\sum_1^\infty t c_t^2 \right]^2 < \infty \text{ under } (\mathcal{S}_5). \end{aligned}$$

For the first term, consider

$$\begin{aligned} R_N &= \sum_{s=1}^N \left[\sum_{t=1}^s c_t c_{t+s} \right]^2 \leq \sum_{s=1}^N (\sum_{t=1}^s t^2 c_t^2) (\sum_{t=1}^s c_{t+s}^2 t^{-2}) \\ &\leq (\sum_1^\infty t c_t^2) \sum_{s=1}^N s^2 \sum_{t=1}^s c_{t+s}^2 t^{-2} = (\sum_1^\infty t c_t^2) \sum_{t=1}^N t^{-2} \sum_{s=t}^N s^2 c_{t+s}^2 \\ &\leq (\sum_1^\infty t c_t^2) \sum_{t=1}^N t^{-1} \sum_{r=2t}^{N+t} r^2 c_r^2 \leq (\sum_1^\infty t c_t^2) \sum_{t=1}^N t^{-2} \sum_{r=t}^{2N} r^2 c_r^2 \\ &\leq (\sum_1^\infty t c_t^2) \sum_{t=1}^{2N} t^{-2} \sum_{r=t}^{2N} r^2 c_r^2 = (\sum_1^\infty t c_t^2) \sum_{r=1}^{2N} r^2 c_r^2 \sum_{t=1}^r t^{-2} \\ &\leq (\sum_1^\infty t c_t^2) \sum_{r=1}^{2N} r c_r^2 \\ &\leq \left[\sum_1^\infty t c_t^2 \right]^2. \end{aligned}$$

Since R_N is nondecreasing, $R_N \rightarrow R_\infty < \infty$. Thus, the first term is also finite if (\mathcal{S}_5) holds, thereby proving the lemma. \square

5.11. PROOF OF THEOREM 3.9. Set $n^{-1}w_n = n^{-1}\Sigma_1^n[X_t - C(1)\epsilon_t]$. We need to show

$$(52) \quad \Sigma_1^n^{-2}E(w_n^2) < \infty$$

so that $n^{-1}w_n \xrightarrow{\text{a.s.}} 0$ and thus $n^{-1}\Sigma_1^n X_t \xrightarrow{\text{a.s.}} 0$ since $n^{-1}\Sigma_1^n \epsilon_t \xrightarrow{\text{a.s.}} 0$ under (\mathcal{A}_2) . Now

$$E(\Sigma_1^n X_t \Sigma_1^n \epsilon_s) = \Sigma_1^n E(X_t \Sigma_1^t \epsilon_s) = \Sigma_1^n \Sigma_1^t c_{t-s} \sigma_\epsilon^2 = \Sigma_1^n (\Sigma_0^{t-1} c_j) \sigma_\epsilon^2 = \sigma_\epsilon^2 \Sigma_1^n [C(1) - \tilde{c}_{t-1}]$$

so that

$$\begin{aligned} n^{-2}E(w_n^2) &= E(\bar{X}^2) + n^{-1}C(1)^2 \sigma_\epsilon^2 - 2n^{-2}C(1) \sigma_\epsilon^2 \Sigma_1^n [C(1) - \tilde{c}_{t-1}] \\ &= [E(\bar{X}^2) - n^{-1} \sigma_\epsilon^2 C(1)^2] + 2\sigma_\epsilon^2 C(1) [n^{-2} \Sigma_0^{n-1} \tilde{c}_t] \\ (53) \quad &= a_n + 2\sigma_\epsilon^2 C(1) b_n, \text{ say.} \end{aligned}$$

Next

$$\begin{aligned} \Sigma_1^n |b_n| &\leq \Sigma_1^n n^{-2} \Sigma_0^{n-1} |\tilde{c}_t| \leq \Sigma_0^n |\tilde{c}_t| \Sigma_{t+1}^n n^{-2} \\ &\leq \Sigma_1^n |\tilde{c}_t| [\Sigma_{t+1}^n [n(n-1)]^{-1}] + |\tilde{c}_0| \Sigma_1^n n^{-2} \\ &\leq \Sigma_1^n |\tilde{c}_t| t^{-1} + |\tilde{c}_0| \Sigma_1^n n^{-2} \\ &\leq \Sigma_1^n t^{-1} \Sigma_{t+1}^n |c_s| + |\tilde{c}_0| \Sigma_1^n n^{-2} \\ &= \Sigma_1^n |c_s| \Sigma_1^{s-1} t^{-1} + |\tilde{c}_0| \Sigma_1^n n^{-2} \\ &\leq \Sigma_1^n |c_s| (1 + \ell n s) + (\Sigma_1^n |c_s|) (\Sigma_1^n n^{-2}) \end{aligned}$$

by Lemma 5.1(c). Further

$$\begin{aligned} a_n &= n^{-1}(\gamma_0 + 2\Sigma_1^{n-1} \gamma_r) - 2n^{-2} \Sigma_1^{n-1} r \gamma_r - n^{-1}(\gamma_0 + 2\Sigma_1^n \gamma_r) \\ (54) \quad &= -2n^{-1} \Sigma_1^n \gamma_r - 2n^{-2} \Sigma_1^{n-1} r \gamma_r \end{aligned}$$

and, then, by use of Lemma 5.1 again we deduce that:

$$\begin{aligned}
\Sigma_1^\infty |a_n| &\leq 2\Sigma_1^\infty n^{-1} \Sigma_n^\infty |\gamma_r| + 2\Sigma_1^\infty n^{-2} \Sigma_1^{n-1} |\gamma_r| \\
&\leq 2\Sigma_1^\infty |\gamma_r| (\ln r + 1) + 2\Sigma_1^\infty |\gamma_r| \Sigma_{n=r+1}^\infty n^{-2} \\
&\leq 2\Sigma_1^\infty |\gamma_r| (\ln r + 1) + 2\Sigma_1^\infty |\gamma_r|.
\end{aligned}$$

But

$$\begin{aligned}
\Sigma_1^\infty \ln r |\gamma_r| &\leq \Sigma_{r=1}^\infty \Sigma_{s=0}^\infty \ln r |c_s| |c_{s+r}| \\
&\leq \Sigma_{s=0}^\infty |c_s| \Sigma_{r=1}^\infty \ln(r+s) |c_{s+r}| \\
&= \Sigma_{s=0}^\infty |c_s| \Sigma_{p=s+1}^\infty \ln p |c_p| \\
&\leq \Sigma_{p=1}^\infty \ln p |c_p| \Sigma_{s=0}^{p-1} |c_s| \\
&\leq \Sigma_{p=1}^\infty \ln p |c_p| \Sigma_{s=0}^\infty |c_s| (\ln s + 1) \\
&< \infty
\end{aligned}$$

under (S_6) . Thus $\Sigma_1^\infty |a_n| < \infty$, $\Sigma_1^\infty |b_n| < \infty$ and (52) follows. \square

5.12. PROOF OF THEOREM 3.10. From (21) and (54) we have

$$n^{-1} E(w_n^2) = -2\Sigma_n^\infty \gamma_r - 2n^{-1} \Sigma_1^{n-1} r \gamma_r + 2\sigma_\epsilon^2 C(1) n^{-1} \Sigma_0^{n-1} \bar{c}_t.$$

But $\Sigma_1^n \gamma_r$ converges to $\Sigma_1^\infty \gamma_r$, which is finite under (S_7) since $\gamma_0 + 2\Sigma_1^\infty \gamma_r = C(1)^2 \sigma_\epsilon^2$. Thus, $\Sigma_n^\infty \gamma_r \rightarrow 0$ as $n \rightarrow \infty$. Further, by Kronecker's lemma $n^{-1} \Sigma_1^n r \gamma_r \rightarrow 0$. Finally, if $\bar{c}_t \rightarrow 0$ as $t \rightarrow \infty$ then by the Toeplitz lemma $n^{-1} \Sigma_1^n \bar{c}_t \rightarrow 0$. However, $\bar{c}_t = \Sigma_{t+1}^\infty c_s \rightarrow 0$ as $t \rightarrow \infty$ because $\Sigma_0^\infty c_s = C(1)$ is convergent under (S_7) . Hence $n^{-1/2} w_n \rightarrow_p 0$ and $n^{-1/2} \Sigma_1^n X_t = C(1) n^{-1/2} \Sigma_1^n \epsilon_t + o_p(1) \rightarrow_d N(0, \sigma_\epsilon^2 C(1)^2)$ as required. \square

5.13. PROOF OF THEOREM 3.12. This follows the proof of Theorem 3.1. By Remark 2.9(ii) $n^{-1} \Sigma_1^n \epsilon_t \rightarrow_{a.s.} 0$ under (B_2) . The result then follows if $E(\bar{\epsilon}_n^2)$ is bounded above uniformly in n , which it is in view of (S_1) and (B_2) . \square

5.14. PROOF OF THEOREM 3.13. This follows the proof of Theorem 3.2. We just note that theorem LLN ensures that $n^{-1}\sum_1^n \epsilon_t \xrightarrow{\text{a.s.}} 0$ while (\mathcal{B}_1) ensures that $E|\epsilon_t| \leq E|Z| < \infty$ so that $\sum_1^{\infty} \bar{c}_t |\epsilon_t| < \infty$ a.s. \square

5.15. PROOF OF THEOREM 3.14. The proof is similar to that of Theorem 3.3. We work from equation (18). The CLT for $n^{-1/2}\sum_1^n \epsilon_t$ follows from theorem CLT if

$$(55) \quad n^{-1}\sum_1^n \epsilon_t^2 \xrightarrow{\text{p}} \sigma_\epsilon^2$$

and

$$(56) \quad n^{-1}\sum_1^n E[\epsilon_t^2 1(\epsilon_t^2 > n\epsilon)] \rightarrow 0.$$

Now (55) follows from (\mathcal{B}_2) , as indicated in Remark 2.9(ii). Next from Billingsley (1968, p. 223) we have

$$\begin{aligned} E[\epsilon_t^2 1(\epsilon_t^2 > n\epsilon)] &= \epsilon n P(\epsilon_t^2 > n\epsilon) + \int_{n\epsilon}^{\infty} P(\epsilon_t^2 > s) ds \\ &\leq c[\epsilon n P(Z^2 > n\epsilon) + \int_{n\epsilon}^{\infty} P(Z^2 > s) ds] \\ &= cE[Z^2 1(Z^2 > n\epsilon)]. \end{aligned}$$

Thus, dominated convergence and $E(Z^2) < \infty$ yield (56). This establishes that $n^{-1/2}\sum_1^{[nr]} \epsilon_t \xrightarrow{\text{d}} \sigma_\epsilon W(r)$. Theorem 3.14(a) now holds because $n^{-1/2}\tilde{\epsilon}_0, n^{-1/2}\tilde{\epsilon}_n \xrightarrow{\text{p}} 0$ since

$$E(n^{-1}\tilde{\epsilon}_n^2) = n^{-1}\sum_0^{\infty} \tilde{c}_j^2 E(\epsilon_{n-j}^2) \leq n^{-1}\sum_0^{\infty} \tilde{c}_j^2 E(Z^2) \rightarrow 0,$$

under (\mathcal{S}_1) and (\mathcal{B}_2) .

To prove the IP (part (b)) we need to verify (19) or equivalently (20), i.e.

$$n^{-1}\sum_1^n [\tilde{\epsilon}_k^2 1(\tilde{\epsilon}_k^2 > n\delta)] \xrightarrow{\text{p}} 0, \text{ for any } \delta > 0.$$

This holds if

$$n^{-1} \sum_1^n E[\tilde{\epsilon}_k^2 1(\tilde{\epsilon}_k^2 > n\delta)] \rightarrow 0$$

which holds if the variables $\tilde{\epsilon}_k^2$ are uniformly integrable, which holds by Minkowski's inequality applied to $E|\tilde{\epsilon}_k|^{2+\eta}$ and (S_3) when $\sup_t E|\epsilon_t|^{2+\eta} < \infty$ for some $\eta > 0$. Thus, (S_3) and the strengthened version of (B_2) suffice to establish the IP. \square

5.16. PROOF OF THEOREM 3.15. This follows the proof of Theorem 3.6. We need to verify (41)–(44). As noted in Remark 2.9(ii) (B_2) gives (44). For (43) the same proof as in 5.7 applies and we just need

$$E(\sum_1^{\infty} \tilde{f}_{0k} \epsilon_{-k}^2) < \infty.$$

But $\sum_1^{\infty} \tilde{f}_{0k} < \infty$ under (S_5) , and by (B_2) we have

$$E(\epsilon_{-k}^2) = \int_0^{\infty} P(\epsilon_{-k}^2 > s) ds \leq c \int_0^{\infty} P(Z^2 > s) ds = cE(Z^2) < \infty.$$

For (41) the same proof as in 5.7 again applies. All we need is the result of Lemma 5.8 and this holds under (S_5) and the strengthened version of (B_2) since $\sup_t E(\epsilon_t^4) \leq E(Z^4) < \infty$. Finally, we consider (42). As in §5.7, $T_n = \sum_1^n t^{-1} \epsilon_t^f \epsilon_{t-1}^f$ is still a martingale and we find

$$\begin{aligned} E(T_n^2) &= \sum_1^n t^{-2} E[\epsilon_t^2 (\epsilon_{t-1}^f)^2] = \sum_1^n t^{-2} \sum_1^{\infty} \gamma_j^2 E(\epsilon_t^2 \epsilon_{t-j}^2) \\ &\leq \sup_t E(\epsilon_t^4) \sum_1^n t^{-2} \sum_1^{\infty} \gamma_j^2 < \infty. \quad \square \end{aligned}$$

5.17. PROOF OF THEOREM 3.17. From (29) we have

$$(57) \quad n^{1/2}(r_h - \rho_h) = \left[n^{-1} \sum_1^n X_t^2 \right]^{-1} n^{-1/2} \sum_1^n \epsilon_t \epsilon_{t-1}^d + o_p(1)$$

where $\epsilon_{t-1}^d = \sum_{r=1}^{\infty} g_r \epsilon_{t-r}$, $g_r = f_{h+r}(1) + f_{h-r}(1) - \rho_h(f_r(1) + f_{-r}(1))$. We shall work first with the finite sum $\epsilon_{t-1}^{d\ell} = \sum_{r=1}^{\ell} g_r \epsilon_{t-r}$ and apply theorem CLT to $n^{-1/2} \sum_1^n \epsilon_t \epsilon_{t-1}^{d\ell}$.

We need to verify

$$(58) \quad n^{-1} \sum_1^n Z_{t\ell}^2 \rightarrow_p x_\ell(h)$$

$$(59) \quad n^{-1} \sum_1^n \mathbb{E}[Z_{t\ell}^2 \mathbf{1}(Z_{t\ell}^2 > \epsilon n)] \rightarrow 0$$

where $Z_{t\ell} = \epsilon_t \epsilon_{t-1}^{\text{d}\ell}$ and $x_\ell(h) = \sigma^4 \sum_{r,s=1}^\ell \tau_{rs} g_r g_s$. But (58) holds since by theorem LLN

$$(60) \quad n^{-1} \sum_1^n [\epsilon_t^2 (\epsilon_{t-1}^{\text{d}})^2 - \mathbb{E}(\epsilon_t^2 | \mathcal{F}_{t-1}) (\epsilon_{t-1}^{\text{d}})^2] \rightarrow_p 0.$$

Observe that the conditions of theorem LLN hold because a suitable constant multiple of Z^4 is a bounding r.v. for $Z_{t\ell}^2 = \epsilon_t^2 \left[\epsilon_{t-1}^{\text{d}\ell} \right]^2$. In particular, for any $u \geq 0$ we have:

$$\mathbb{P} \left[\epsilon_t^2 (\epsilon_{t-1}^{\text{d}\ell})^2 > u \right] \leq \mathbb{P} \left[\epsilon_t^4 + (\epsilon_{t-1}^{\text{d}\ell})^4 > 2u \right] \leq \mathbb{P}(\epsilon_t^4 > u) + \mathbb{P} \left[(\epsilon_{t-1}^{\text{d}\ell})^4 > u \right].$$

Now Z^4 is a bounding r.v. for ϵ_t^4 so it remains to find a suitable bounding r.v. for $(\epsilon_{t-1}^{\text{d}\ell})^4$. But

$$\begin{aligned} \mathbb{P}(|\epsilon_{t-1}^{\text{d}\ell}| > u) &\leq \mathbb{P}(\Sigma_1^\ell |g_r| |\epsilon_{t-r}| > u) \\ &\leq \mathbb{P}(U_1^\ell \{ |g_r| |\epsilon_{t-r}| > u/\ell \}) \\ &\leq \Sigma_1^\ell \mathbb{P}(|\epsilon_{t-r}| > u/\ell g), \quad g = \Sigma_1^\infty |g_r| \\ &\leq c\ell \mathbb{P}(|Z| > u/\ell g) \end{aligned}$$

and, under the strengthened (B_2) , $(\ell g Z)^4$ is a bounding r.v. for $(\epsilon_{t-1}^{\text{d}\ell})^4$ and hence $\epsilon_t^2 (\epsilon_{t-1}^{\text{d}\ell})^2$. Observe that $g = \Sigma_1^\infty |g_r| < \infty$ since $|g_r| \leq |f_{h+r}(1)| + |f_{h-r}(1)| + |\rho_h| (|f_r(1)| + |f_{-r}(1)|)$ and e.g.

$$\Sigma_{r=1}^\infty |\Sigma_{k=0}^\infty c_k c_{k+h+r}| \leq \Sigma_{k=0}^\infty |c_k| \Sigma_{s=0}^\infty |c_s| < \infty,$$

under (S_8) . Also, under the strengthened (B_2) we have

$$n^{-1} \sum_1^n [\mathbb{E}(\epsilon_t^2 | \mathcal{F}_{t-1}) (\epsilon_{t-1}^{\text{d}\ell})^2] \rightarrow_{\text{a.s.}} \sigma^4 \sum_{r,s=1}^\ell \tau_{rs} g_r g_s = x_\ell(h).$$

Combining this with (60) we get (58) as required. Since $\mathbb{E}(Z^4) < \infty$, (59) also holds and

thus we have

$$n^{-1/2} \sum_1^n \epsilon_t \epsilon_{t-1}^{\text{d}\ell} \rightarrow_d Y_\ell =_d N(0, x_\ell(h)).$$

Next, as $\ell \rightarrow \infty$ we have

$$Y_\ell \rightarrow_d N(0, x_\infty(h))$$

where $x_\infty(h) = \sigma^4 \sum_{r,s=1}^\infty \tau_{rs} g_r g_s$, which is absolutely convergent because τ_{rs} is uniformly bounded and $\sum_1^\infty |g_r| < \infty$. Further, setting

$$\epsilon_{t-1}^\ell = \epsilon_{t-1}^{\text{d}} - \epsilon_{t-1}^{\text{d}\ell} = \sum_{r=\ell+1}^\infty g_r \epsilon_{t-r}$$

we have for any $\eta > 0$

$$P(|n^{-1/2} \sum_1^n \epsilon_t \epsilon_{t-1}^\ell| > \eta) < \eta^{-2} n^{-1} \sum_1^n E[\epsilon_t^2 (\epsilon_{t-1}^\ell)^2] = \eta^{-2} \sum_{r,s=\ell+1}^\infty \tau_{rs} g_r g_s.$$

Thus

$$\begin{aligned} \lim_{\ell \rightarrow \infty} \limsup_{n \rightarrow \infty} P(|n^{-1/2} \sum_1^n \epsilon_t \epsilon_{t-1}^\ell| > \eta) \\ < \lim_{\ell \rightarrow \infty} \eta^{-2} \sum_{r,s=\ell+1}^\infty \tau_{rs} g_r g_s = 0 \end{aligned}$$

and it follows directly that

$$n^{-1/2} \sum_1^n \epsilon_t \epsilon_{t-1}^{\text{d}} \rightarrow_d N(0, x_\infty(h))$$

(cf. Billingsley (1968), theorem 4.2, p. 25). Finally, we note that $n^{-1} \sum_1^n X_t^2 \rightarrow_{\text{a.s.}} \gamma_0$ by Theorem 3.15, so that from (57)

$$n^{1/2} (r_h - \rho_h) \rightarrow_d N(0, \gamma_0^{-2} \sigma^4 \sum_{r,s=1}^\infty \tau_{rs} g_r g_s) = N(0, x(h))$$

since $\gamma_0^{-1} g_r = \rho_{h+r} + \rho_{h-r} - \rho_h (\rho_r + \rho_{-r}) = \rho_{h+r} + \rho_{h-r} - 2\rho_h \rho_r$. \square

5.18. PROOF OF THEOREM 3.19. From (15)

$$a_n^{-1} \Sigma_1^n X_t = C(1) a_n^{-1} \Sigma_1^n \epsilon_t + a_n^{-1} (\tilde{\epsilon}_0 - \tilde{\epsilon}_n)$$

and $a_n^{-1} \Sigma_1^n \epsilon_t \rightarrow_d U_\alpha(1)$ under (C_1) by (10) so the result follows provided

$$a_n^{-1} (\tilde{\epsilon}_0 - \tilde{\epsilon}_n) \rightarrow_p 0.$$

But under (C_1) and (S_9) , $\tilde{\epsilon}_n$ is strictly stationary and in $\mathcal{D}(\alpha)$ (e.g. Brockwell and Davis (1987), p. 481). Thus for any $\delta > 0$ we have

$$P(a_n^{-1} |\tilde{\epsilon}_n| > \delta) = P(|\tilde{\epsilon}_n| > a_n \delta) = O(n^{-1})$$

and $a_n^{-1} |\tilde{\epsilon}_n| \rightarrow_p 0$, as required. \square

5.19. LEMMA

(a) Under (S_9) , $\Sigma_{j=0}^\infty |\tilde{f}_{kj}|^{p/2} < \infty$ and $\Sigma_{r=1}^\infty |\Sigma_{k=0}^\infty \tilde{f}_{h+r,k}|^p < \infty$.

(b) Under (S_{10}) , $\Sigma_{r=-\infty, r \neq 0}^\infty |f_{h+r}(1)|^p < \infty$.

PROOF

$$\begin{aligned} \Sigma_{j=0}^\infty |\tilde{f}_{kj}|^{p/2} &= \Sigma_{j=0}^\infty |\Sigma_{s=j+1}^\infty c_s c_{s+k}|^{p/2} \leq \Sigma_{j=0}^\infty \Sigma_{s=j+1}^\infty |c_s|^{p/2} |c_{s+k}|^{p/2} \\ &= \Sigma_{s=1}^\infty |c_s|^{p/2} |c_{s+k}|^{p/2} \Sigma_{j=0}^{s-1} 1 = \Sigma_{s=1}^\infty s |c_s|^{p/2} |c_{s+k}|^{p/2} \\ &= \Sigma_{s=1}^\infty s^{1/2} |c_s|^{p/2} s^{1/2} |c_{s+k}|^{p/2} < \left[\Sigma_{s=1}^\infty s |c_s|^p \right]^{1/2} \left[\Sigma_{s=1}^\infty (s+k) |c_{s+k}|^p \right]^{1/2} \\ &< \Sigma_{s=1}^\infty s |c_s|^p < \infty \end{aligned}$$

under (S_9) , giving (a). Use Σ'_r to signify the summation $\Sigma_{r=-\infty, r \neq 0}^\infty$ and we have

$$\begin{aligned} \Sigma'_r |f_{h+r}(1)|^p &= \Sigma'_r |\Sigma_{s=0}^\infty c_s c_{s+h+r}|^p \leq \Sigma'_r \Sigma_0^\infty |c_s|^p |c_{s+h+r}|^p \\ &= \Sigma_{s=0}^\infty |c_s|^p \Sigma'_r |c_{s+h+r}|^p < \left[\Sigma_0^\infty |c_s|^p \right]^2 < \infty \end{aligned}$$

under (S_{10}) . \square

5.20. PROOF OF THEOREM 3.19. We use (28) and write

$$(61) \quad X_t X_{t+h} = f_h(1) \epsilon_t^2 + \epsilon_{t+h} \epsilon_t^h - (1-L) \tilde{f}_h(L) \epsilon_t^2 \\ - (1-L) \sum_{r=1}^{\infty} [\tilde{f}_{h+r}(L) \epsilon_{t+h-r} \epsilon_{t+h} + \tilde{f}_{h-r}(L) \epsilon_{t+h+r} \epsilon_{t+h}]$$

where $\epsilon_t^h = \sum_{r=1}^{\infty} [f_{h+r}(1) \epsilon_{t+h-r} + f_{h-r}(1) \epsilon_{t+h+r}]$. In view of Lemma 5.19(b) $\epsilon_t^h \in \mathcal{D}(\alpha)$ and, since ϵ_t is independent of ϵ_t^h , we have $\epsilon_{t+h} \epsilon_t^h \in \mathcal{D}(\alpha)$. Thus,

$$(62) \quad a_n^{-2} \sum_1^n \epsilon_{t+h} \epsilon_t^h \rightarrow_p 0.$$

Next set $\tilde{\epsilon}_n^2 = \tilde{f}_n(L) \epsilon_n^2 = \sum_{j=0}^{\infty} \tilde{f}_{hj} \epsilon_{n-j}^2$. By Lemma 5.19(a) the series for $\tilde{\epsilon}_n^2$ converges a.s. and $\tilde{\epsilon}_n^2 \in \mathcal{D}(\alpha/2)$. Hence for any $\delta > 0$ $P(a_n^{-2} \tilde{\epsilon}_n^2 > \delta) = O(n^{-1})$ and so

$$(63) \quad a_n^{-2} \tilde{\epsilon}_n^2 \rightarrow_p 0.$$

Next note that under (\mathcal{S}_9)

$$\sum_{r=1}^{\infty} \sum_{k=0}^{\infty} |\tilde{f}_{h+r,k}|^p \leq \sum_{r=1}^{\infty} \sum_{k=0}^{\infty} \sum_{s=k+1}^{\infty} |c_s c_{s+h+r}|^p \\ = \sum_{r=1}^{\infty} \sum_{s=1}^{\infty} s |c_s|^p |c_{s+h+r}|^p < (\sum_{s=1}^{\infty} s |c_s|^p) (\sum_{r=0}^{\infty} |c_s|^p) < \infty,$$

so that $\sum_{r=1}^{\infty} \sum_{k=0}^{\infty} \tilde{f}_{h+r,k} \epsilon_{t+h-r-k} \epsilon_{t+h-k}$ converges a.s. and is in $\mathcal{D}(\alpha)$. Thus

$$(64) \quad a_n^{-2} \sum_{r=1}^{\infty} \tilde{f}_{h+r}(L) \epsilon_{n+h-r} \epsilon_{n+h} \rightarrow_p 0$$

and similarly

$$(65) \quad a_n^{-2} \sum_{r=1}^{\infty} \tilde{f}_{h-r} \epsilon_{n+h+r} \epsilon_{n+h} \rightarrow_p 0.$$

Summing (61), scaling by a_n^{-2} and using (62)–(65) we deduce that

$$(66) \quad a_n^{-2} \sum_1^n X_t X_{t+h} = f_h(1) a_n^{-2} \sum_1^n \epsilon_t^2 + o_p(1) \\ \rightarrow_d f_h(1) \int_0^1 (dU_{\alpha})^2$$

by (12). The joint convergence of $a_n^{-2}[\Sigma_1^n X_t^2, \dots, \Sigma_1^n X_t X_{t+h}]$ follows directly. Part (b) follows since

$$r_h = \left[\Sigma_1^n X_t^2 \right]^{-1} (\Sigma_1^n X_t X_{t+h}) = f_h(1)/f_h(0) + o_p(1). \quad \square$$

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